

# Adaptive Kalman Filtering

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The increased power of small computers makes the use of parameter estimation methods attractive. Such methods have a number of uses in analytical chemistry. When valid models are available, many methods work well, but when models used in the estimation are in error, most methods fail. Methods based on the Kalman filter, a linear recursive estimator, may be modified to perform parameter estimation with erroneous models. Modifications to the filter involve allowing the filter to adapt the measurement model to the experimental data through matching the theoretical and observed covariance of the filter innovations sequence. The adaptive filtering methods that result have a number of applications in analytical chemistry.

**Key words:** automated covariance estimation; Kalman filter; multicomponent analysis.

## 1. Introduction

The increased computational power available from small computers has prompted a re-evaluation of the methods used in reducing data obtained from a chemical analysis. Many of the responses obtained from chemical analyses are suited to mathematical analysis by methods which estimate the parameters that generate the response; these parameters are generally concentrations. For parameter estimation to be successful, an accurate model of the behavior of the chemical system is necessary. The model used need not be theoretical; empirical models based on experimental results or on a numerical simulation of the chemical system are often satisfactory

as well. When valid models are available, the parameters associated with the model may be obtained with a variety of methods. Some that have seen extensive use in analytical chemistry include analysis of the chemical data using linear least squares [1], nonlinear least squares analysis [2,3], and Kalman filtering [4-6].<sup>1</sup>

The methods mentioned above all work well with accurate models, but are much less satisfactory when used with models containing errors that can arise from many sources. Theoretical models, or models based on simulation, may not describe the physics or chemistry of a system well enough to predict system responses to the accuracy desired. Small changes in the experimental conditions used for data acquisition may perturb experimentally-obtained models, leading to errors when these models are used to analyze subsequent experiments. And, it may be impossible, because of the effects of chemical equilibria, to obtain independent responses for some of the chemical species included in a model for a complex system, leading to "chemical" model errors.

Relatively few methods have been developed to compensate for model errors affecting multicomponent quantitation. Approaches using factor analysis [7] have

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<sup>1</sup> Bracketed figures indicate literature references.

been developed for situations where the model is unknown, but these approaches are generally limited to very few components [8], and it is difficult to incorporate additional *a priori* information into such methods. An alternative approach has used the Kalman filter. The Kalman filter is a linear, recursive estimator which yields optimal estimates for parameters associated with a valid model [9,10]. Several methods, classified under the term "adaptive filtering," have been developed to permit the filter to produce accurate parameter estimates in the presence of model errors [11–15]. This paper summarizes the development of an adaptive Kalman filter for use in the mathematical analysis of overlapped multicomponent chemical responses.

## 2. Theory

**Kalman Filtering.** The Kalman filter has received some attention for the analysis of multicomponent chemical responses [4,6,16,17]. Because most models relating chemical responses to concentrations are linear, application of the Kalman filter is straightforward. The filter model is comprised of two equations. The system model, which describes the time evolution of the desired parameters, is, in state-space notation

$$\mathbf{X}(k) = \mathbf{F}(k, k-1)\mathbf{X}(k-1) + \mathbf{w}(k) \quad (2.1)$$

where  $\mathbf{X}$  is a  $n \times 1$  column vector of state variables describing the chemical system, where  $\mathbf{F}$  is an  $n \times n$  matrix describing how the states change with time,  $\mathbf{w}$  is a vector describing noise contributions to the system model, and where  $k$  indicates time or some other independent variable which meets the noise requirements given below. For state-invariant systems,  $\mathbf{F}$  reduces to the identity matrix  $\mathbf{I}$ . Because multicomponent analysis is most often performed under conditions where concentrations are constant over the time frame involved, the case where  $\mathbf{X}$  is time-invariant is considered here.

The second equation describes the measurement process by relating the measured response  $z(k)$ , to the filter states. For a single sensor, the measurement model is given by

$$z(k) = \mathbf{H}^T(k)\mathbf{X}(k) + v(k) \quad (2.2)$$

where  $\mathbf{H}^T(k)$  is a  $1 \times n$  vector relating the response at point  $k$  to the  $n$  states, and the scalar  $v(k)$  is the noise contribution of the measurement process. For example, in absorption spectrophotometry,  $z(k)$  is an absorbance measurement at some wavelength  $k$ , and  $\mathbf{H}^T(k)$  is the vector of absorption coefficients at that wavelength for all chemical species included in the model. The mea-

surement model is easily extended to systems with multiple sensors.

The two noise processes in the Kalman filter,  $\mathbf{w}(k)$  and  $v(k)$ , are usually assumed to be independent, zero-mean, white noise processes. The matrix  $\mathbf{Q}(k)$ , defined as the covariance of the noise in the system model, is taken as approximately zero for the time invariant system discussed in this paper. The scalar quantity  $R(k)$  is the variance of the noise in the measurement process.

The Potter-Schmidt square-root algorithm, one implementation of the Kalman filter [18], is given in table 1. The details of this algorithm have been discussed elsewhere [18,19]. Initial guesses for the filter states and for the covariance matrix  $\mathbf{P}$  are required to start the filter. Estimates of  $\mathbf{X}$  and  $\mathbf{P}$  depend on  $k$ , and because both are projected ahead of the data (in eqs 2.3 and 2.4) by the filter, the notation  $(j|k)$  is used to indicate that the estimate is made at point  $j$ , based on data obtained up through point  $k$ . The filter output consists of estimates  $\hat{\mathbf{X}}$ , as well as  $\hat{\mathbf{P}}$ . In analytical chemistry, these are often estimates of concentrations and of the error in the concentrations.

Table 1. Algorithm equations for the square root Kalman filter.

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State estimate extrapolation	
$\mathbf{X}(k k-1) = \mathbf{F}(k k-1)\mathbf{X}(k-1 k-1)$	(2.3)
Covariance square root extrapolation	
$\mathbf{S}(k k-1) = \mathbf{F}(k, k-1)\mathbf{S}(k-1 k-1)\mathbf{F}^T(k, k-1)$	(2.4)
where	
$\mathbf{F}(k, k-1) = \mathbf{I}$	
$\mathbf{P} = \mathbf{S}\mathbf{S}^T$	
Kalman gain:	
$\mathbf{K}(k) = \mathbf{a}\mathbf{S}(k k-1)\mathbf{G}(k)$	(2.5)
where	
$\mathbf{G}(k) = \mathbf{S}^T(k k-1)\mathbf{H}(k)$	(2.6)
$1/\mathbf{a} = \mathbf{G}^T(k)\mathbf{G}(k) + R(k)$	(2.7)
$\mathbf{a} = (1 + (\mathbf{a} \cdot R(k))^{1/2})^{-1}$	(2.8)
State estimate update:	
$\mathbf{X}(k k) = \mathbf{X}(k k-1) + \mathbf{K}(k)[z(k) - \mathbf{H}^T(k)\mathbf{X}(k k-1)]$	(2.9)
Covariance square root update:	
$\mathbf{S}(k k) = \mathbf{S}(k k-1) - \mathbf{a}\mathbf{d}\mathbf{S}(k k-1)\mathbf{G}(k)\mathbf{G}^T(k)$	(2.10)
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**Adaptive Kalman Filtering.** Errors can occur in both of the models used in the Kalman filter. Errors in the system model arise if the system was taken as time-invariant, but was actually composed of time-dependent states. Errors in the measurement model arise from underestimating the number of components involved in

the state vector (which can be thought of as incorrectly setting values in  $\mathbf{H}^T(k)$  to zero for one of the possible elements in the state vector), or by use of inaccurate values in  $\mathbf{H}^T(k)$ . Either type of error produces a sub-optimal filter, in that the accuracy of the filter's estimates are severely degraded. Many methods for compensating these model errors make use of the filter innovations sequence,  $\nu(k)$ , defined as

$$\nu(k) = z(k) - \mathbf{H}^T(k)\hat{\mathbf{X}}(k|k-1). \quad (2.11)$$

The innovations sequence can be used to construct a measure of the optimality of the filter; a necessary and sufficient condition for an optimal filter is that this sequence be a white noise process [10]. An optimal filter is one that minimizes the mean square estimation error  $E\{(\mathbf{X} - \hat{\mathbf{X}})(\mathbf{X} - \hat{\mathbf{X}})^T\}$ . A suboptimal filter may generate results which show large estimation errors, or even a divergence of the errors [11]. The aim of an adaptive filter is to reduce or bound these errors by modifying, or adapting, the models used in the Kalman filter to the real data.

Several methods for controlling error divergence in the filter have been reported [11-15]. Most involve cases where  $Q$  is poorly known, the situation which arises when the time-dependence of states is incorrectly modeled. These include methods based on Bayesian estimation and maximum likelihood estimation [14], correlation methods [14], and covariance matching techniques [14,15]. The last method has also been suggested for use when  $Q$  is known, but  $R$  is unknown, the situation that arises when the number of components in the state is underestimated, or when the measurement model is otherwise incorrect. Because errors in the number of components and in the response factors used in the measurement model are common in multicomponent chemical analysis, covariance matching is used to develop the filter discussed here.

The aim of covariance matching is to insure that the residuals remain consistent with the theoretical covariances. The covariance of the innovations sequence  $\nu(k)$  is [14]

$$E[\nu(k) \cdot \nu(k)^T] = \mathbf{H}^T(k)P(k|k-1)\mathbf{H}(k) + R(k). \quad (2.12)$$

If the actual covariance of  $\nu(k)$  is much larger than the covariance obtained from the Kalman filter, either  $Q$  or  $R$  should be increased to prevent divergence. In either case, this has the effect of increasing  $P(k|k-1)$ , thus bringing the actual covariance of  $\nu(k)$  closer to that given in eq 2.12. This also has the effect of decreasing the filter gain matrix,  $K$ , thereby "closing" the filter to new data which would otherwise be incorrectly interpreted because of errors in the measurement model. In

essence, this amounts to "covering" the errors in the model with noise, then estimating the noise variance. The adaptive estimate of  $R$  at the  $k$ th point, when  $Q$  is known, is

$$R(k) = 1/m \left[ \sum_{j=1}^m \nu(k-j) \cdot \nu(k-j) \right] - \mathbf{H}^T(k)S(k|k-1)S^T(k|k-1)\mathbf{H}(k) \quad (2.13)$$

where  $m$  is the width of an empirically chosen rectangular smoothing window for the innovations sequence. The smoothing operation improves the statistical significance of the estimator for  $R(k)$ , as it now depends on many residuals.

Adaptive estimation of  $R$  allows accurate estimates for the states to be obtained, even in the presence of model errors, because only data for which an accurate model is available are used in the filter. A new measurement model can be constructed from the estimated  $R(k)$ , either by augmenting the  $\mathbf{H}^T$  vector, or by correcting any one of its existing elements; choice of correction or augmentation is arbitrary. For augmentation, the equations

$$H_{n+1}^*(k) = b(k)[R(k+m/2)]^{1/2}, \text{ for } b(k) > 0 \quad (2.14)$$

$$H_{n+1}^*(k) = 0, \text{ for } b(k) < 0 \quad (2.15)$$

apply, where  $H_{n+1}^*(k)$  denotes an element which is incorporated in the  $\mathbf{H}^T$  vector. The term  $(k+m/2)$  arises from the lag induced by averaging  $m$  of the squared innovations. The factor  $b(k)$  is defined as

$$b(k) = 1, \text{ for } \sum_{j=1}^m \nu(k-j+m/2)/m > 0 \quad (2.16)$$

$$b(k) = -1, \text{ for } \sum_{j=1}^m \nu(k-j+m/2)/m < 0. \quad (2.17)$$

Equations 2.16 and 2.17 allow determination of the sign of the model error by evaluating the average of the innovations over the range for which  $R$  was calculated. Equation 2.15 reflects the fact that the relation between the chemical response and concentration, given by  $\mathbf{H}^T$ , is generally positive.

For correction of the  $i$ th component of the vector  $\mathbf{H}^T$ , the expressions

$$H_i^*(k) = H_i(k) + b(k)[R(k+m/2)]^{1/2}, \quad \text{for } H_i^*(k) > 0 \quad (2.18)$$

$$H_i^*(k) = 0, \text{ for } H_i^*(k) < 0 \quad (2.19)$$

apply instead of those given in eqs 2.14 and 2.15. In either case, a valid measurement model can be generated from the adaptive estimation of  $R$ .

Two criteria must be met for this adaptive filter to be useful in the mathematical analysis of multicomponent responses. First, the model to be adaptively corrected must already be correct for some of the values of  $k$  where each of the known components of the model has a measureable response. The second requirement is that the adaptive correction must be performed on a single component. For a single sensor,  $R$  is a scalar, and it is not possible to distinguish the different portions belonging to the different components. It is often feasible, however, to treat model errors as a single, unmodeled component without affecting the accuracy of some or all of the estimated quantities. Although it has been observed that the adaptive estimation of  $R$  by covariance matching is not a sufficient condition for obtaining an improved measurement model [10], application of this approach in the mathematical analysis of multicomponent responses has shown that significant model improvement generally occurs in practice [15,20].

**Automation of the Adaptive Filter.** The adaptive filter requires an initial guess of the states and of their covariances, just as in the ordinary filter. The adaptive estimation of  $R$  affects the calculation of  $P$ , however, and it is found that the diagonal elements of  $P$  decrease as  $R$  decreases. Since the size of  $R$  is directly related to the quality of the measurement model, this relation provides a means by which the quality of the final filter estimates can be judged. Once results are obtained with minimum values for the diagonal elements of the estimated  $P$ , the resulting corrected measurement model better describes the experimental data available, judging from the deterministic variances of fitting before and after the model correction. Because the innovations are not white in the presence of model error, the filter results are no longer guaranteed to be optimal, but now depend on the initial guess. Thus, the adaptive filter must be run several times, with different initial guesses,  $X_0$  and  $P_0$ , to locate those best estimates. This process is easily automated, however. Simplex optimization [21–23] can be used to minimize the metric based on the diagonal elements of the covariance matrix

$$Y = \sum_{i=1}^n \log P_{ii} \quad (2.20)$$

as a function of the initial guesses input to the adaptive filter. We have previously demonstrated that the minima in the variance surface  $Y = f(X_0, P_0)$  correspond well to the minima in an error surface defined by the quantities  $(X - \hat{X})$  [24].

### 3. Application in Analytical Chemistry

**Empirical Model Improvement.** Empirical models have been used with the Kalman filter to study the

chemical speciation of metal ions. One study [20] reported the adaptive correction of the visible photoacoustic spectrum of  $\text{Pr(EDTA)}^-$ . This spectrum was obtained from data collected on solutions containing both  $\text{Pr}^{3+}$  and  $\text{Pr(EDTA)}^-$  species. Direct spectroscopic measurement of  $\text{Pr(EDTA)}^-$  is not simple. A similar approach was also used to obtain the spectrum of  $(\text{UO}_2)_3(\text{OH})^+$ , another ion whose spectrum is difficult to observe in the absence of related chemical species [25]. These studies demonstrate the ability of adaptive filtering to correct for "chemical" errors in the measurement model.

Two other studies used adaptive filtering to model the electrochemical response of an equilibrium mixture of  $\text{Cd}^{2+}$  and  $\text{Cd(NTA)}^-$  [20,26]. The adaptively modeled component, attributed to the reduction of  $\text{Cd}^{2+}$  after dissociation of the  $\text{Cd(NTA)}^-$  complex, was corrected [26] from an approximate model based on digital simulation [19]. The stability constant for the  $\text{Cd(NTA)}^-$  species was estimated from the concentrations obtained from the filter. These studies illustrate the correction of "theoretical" errors in the measurement model by adaptive filtering.

The adaptive filter has also been used to correct empirical models for errors which occurred in data acquisition. An example is the correction of models used for the resolution of overlapped electrochemical responses. Resolved peaks are generally needed to obtain estimates of the component concentration. Small changes in experimental conditions, occurring between the time when data are obtained for use in empirical models and the time when the mixtures are measured, change peak positions slightly. The resulting inaccuracy in the model degrades the accuracy of the resolution obtained with the Kalman filter. Adaptive filtering can correct for these type of model errors, resulting in substantially improved concentration estimates from multicomponent electrochemical responses [20].

**Removal of Interferences.** In many multicomponent analyses, substances which interfere with the chemical analysis are often present. Frequently, these species must be chemically separated, because they are not easily removed in the mathematical analysis of the data. Adaptive estimation of these unknown components of the model is an alternative approach. The feasibility of this has been demonstrated [24] in a visible spectrophotometric analysis, where adaptive filtering was used to quantify  $\text{UO}_2^{2+}$ ,  $\text{Ni}^{2+}$ ,  $\text{Co}^{2+}$  and picric acid in the presence of the "unknown" contaminant  $\text{Cu}^{2+}$ . The errors in estimating species concentrations were typically less than 5%. An adaptive estimation of  $\text{Co}^{2+}$  in the presence of "unknown"  $\text{Cu}^{2+}$ ,  $\text{Ni}^{2+}$ ,  $\text{UO}_2^{2+}$  and picric acid, where interferent species responses strongly overlap that for the species of interest, gave an estimation

error of 14% with a five-fold excess of interferent species. This estimation's lower accuracy results from the adaptive filter's response when its model restrictions are not met, a situation which occurs here as a consequence of the severe overlap of the analyte and interferent responses. Even though this result is of lower accuracy than many of the others reported, it is still remarkable. Unlike the other fitting, this result does not rely on the use of a complete model. Using peak resolution based on an ordinary filtering approach, with the same incomplete measurement model, an error of 200–300% is likely.

#### 4. Conclusion

The automated, adaptive estimation of measurement model covariance permits the application of Kalman filtering in chemical systems where models are poorly known. Although results obtained from the adaptive filter are not guaranteed to be optimal by theory, significant improvement in the accuracy of models and estimated parameters is generally possible in practice.

Restrictions are fairly minor: parts of the model must be known well enough to "open" the filter to the data, and only one component of the model may be adaptively corrected at a time. Adaptive filtering should yield results similar to those obtained from factor analysis using target transformation [27], but the adaptive filter requires only one mixture response, while factor analysis requires several.

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