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A Model for Directional Hurricane Wind Speeds

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1 Introduction

Let \mathbf{X} be an \mathbb{R}^d -valued random variable whose coordinates $\{X_i\}$, $i = 1, \dots, d$, denote hurricane wind speeds in d -directions at a site. Independent samples of \mathbf{X} can be viewed as synthetic hurricane wind speeds occurring in different storms. The random vector \mathbf{X} cannot be Gaussian since the sequence of wind speeds recorded in an arbitrary direction $i = 1, \dots, d$ during different storm has 0's so that the marginal distribution of X_i has a finite mass at 0.

Our objectives are to develop (1) a probabilistic model for \mathbf{X} describing hurricane wind speeds in 16 directions at angles $\theta_i = 22.5^\circ i$, $i = 1, \dots, 16$, (2) a method for calibrating the model for \mathbf{X} to records available at a site, and (3) a Monte Carlo algorithm for generating synthetic hurricane speeds over an arbitrary number of years a selected site.

2 Probability law of hurricane wind speed

Consider the special case in which the coordinates of \mathbf{X} are Bernoulli random variables, that is,

$$X_i = \begin{cases} 0, & \text{probability } 1 - p_i \\ 1, & \text{probability } p_i, \end{cases} \quad (1)$$

where $p_i \in (0, 1)$ for $i = 1, \dots, d$. The values 0 and 1 of a coordinate X_i of \mathbf{X} correspond to 0 and non-zero hurricane wind speeds in direction $i = 1, \dots, d$. The average number of 0's and 1's of X_i in n independent trials are $n(1 - p_i)$ and np_i , respectively. We use the model in Eq. 1 to illustrated difficulties related to the complete probabilistic characterization of the hurricane wind vector \mathbf{X} .

If the coordinates of \mathbf{X} are independent, Eq. 1 defines the probability law of \mathbf{X} . If the coordinates of \mathbf{X} are dependent, additional information is needed to specify \mathbf{X} . Let $p_{k_1, \dots, k_d} = P(\cap_{i=1}^d \{X_i = k_i\})$ with $k_1, \dots, k_d \in \{0, 1\}$ denote the probability that (X_1, \dots, X_d) is equal to a particular string (k_1, \dots, k_d) of 0's and 1's. We note that (1) the probabilities $\{p_{k_1, \dots, k_d}\}$, $k_1, \dots, k_d \in \{0, 1\}$, define uniquely the probability law of \mathbf{X} and (2) $p_{k_1, \dots, k_d} = \prod_{i=1}^d P(X_i = k_i)$ if \mathbf{X} has independent coordinates.

The complete characterization of \mathbf{X} involves two types of difficulties. First, the number of probabilities $\{p_{k_1, \dots, k_d}\}$ defining the probability law of \mathbf{X} increases rapidly with d . For example, suppose that $d = 3$. The probability law of \mathbf{X} is completely defined by $2^d = 8$ probabilities $p_{k_1, k_2, k_3} = P(X_1 = k_1, X_2 = k_2, X_3 = k_3)$, $k_1, k_2, k_3 \in \{0, 1\}$, that the vector

(X_1, X_2, X_3) is equal to $(0, 0, 0)$, $(1, 0, 0)$, $(0, 1, 0)$, $(0, 0, 1)$, $(1, 1, 0)$, $(1, 0, 1)$, $(0, 0, 1)$, and $(1, 1, 1)$. The number of probabilities $\{p_{k_1, \dots, k_d}\}$ is 8; 32; 1,024; and 65,536 for $d = 3$; 5; 10; and 16, respectively. Numerical calculations involving 65,536 probabilities are not feasible. Second, the probabilities $\{p_{k_1, \dots, k_d}\}$ need to be estimated from data. Estimates of these probabilities are likely to be unreliable or even impossible for vectors \mathbf{X} with dimension $d = 8$ or larger if based on records of typical length. These considerations demonstrate the need for developing simplified models for \mathbf{X} that are numerically tractable and their parameters can be estimated reliably from data.

3 Translation model for hurricane wind speeds

We propose a translation non-Gaussian model \mathbf{X}_T for the wind speed vector \mathbf{X} , present a method for estimating the probability law of \mathbf{X}_T , and develop a Monte Carlo algorithm for generating samples of \mathbf{X}_T .

3.1 Model definition

Let p_i and F_i denote the probability that the coordinate X_i , $i = 1, \dots, d$, of \mathbf{X} is not 0 and the distribution of the non-zero values of this coordinate, so that

$$\tilde{F}_i(x) = (1 - p_i) 1(x \geq 0) + p_i F_i(x), \quad i = 1, \dots, d, \quad (2)$$

is the distribution of X_i , where $1(A) = 1$ and 0 if statement A is valid and invalid, respectively. We can view X_i as a generalized Bernoulli variable that is 0 with probability $1 - p_i$ and is a random variable following the distribution F_i with probability p_i .

Consider an \mathbb{R}^d -valued random variable \mathbf{X}_T with coordinates $X_{T,i}$ defined by

$$X_{T,i} = \tilde{F}_i^{-1}(G_i), \quad i = 1, \dots, d, \quad (3)$$

where $\mathbf{G} = (G_1, \dots, G_d)$ is a standard \mathbb{R}^d -valued Gaussian variable, that is, $\text{Mean}[G_i] = 0$, $\text{Var}[G_i] = 1$, and $\text{Covariance}[G_i, G_j] = \rho_{ij}$, $i = 1, \dots, d$. We refer to \mathbf{X}_T as the translation model for \mathbf{X} . The model \mathbf{X}_T has the same marginal distributions as \mathbf{X} irrespective of the covariance matrix $\boldsymbol{\rho} = \{\rho_{ij}\}$ of \mathbf{G} since $X_{T,i}$ is 0 with probability $P(\Phi(G_i) \leq 1 - p_i) = P(G_i \leq \Phi^{-1}(1 - p_i)) = 1 - p_i$ and has distribution F_i with the complement of this probability, that is, $P(X_i \neq 0) = p_i$ for all $i = 1, \dots, d$. The dependence between the coordinates of $\mathbf{X}_{T,i}$ is defined by the covariance matrix $\boldsymbol{\rho}$ of \mathbf{G} and the marginal distributions $\{F_i\}$ of \mathbf{X} . The relationship between the correlation structures of \mathbf{G} and \mathbf{X}_T is discussed in [1] (Section 3.1.1).

The translation model in Eq. 3 has two notable features. The model (1) has, as already stated, the same marginal distributions as \mathbf{X} and (2) is sufficiently simple to be used in applications. A limitation of the model is that the complex dependence between the coordinates of \mathbf{X} is represented approximately.

3.1.1 Parameter estimation

Let $(\mathbf{x}_1, \dots, \mathbf{x}_n)$ be n independent samples of \mathbf{X} , and let $(x_{i,1}, \dots, x_{i,n})$ denote the corresponding n samples of coordinate X_i , $i = 1, \dots, d$. Denote by $(y_{i,1}, \dots, y_{i,m_i})$, $m_i \leq n$,

the sequence of non-zero readings extracted from $(x_{i,1}, \dots, x_{i,n})$. For example, $x_{i,1}$ is not included in $(y_{i,1}, \dots, y_{i,m_i})$ if 0 and $y_{i,1} = x_{i,1}$ if $x_{i,1} \neq 0$.

The probabilities p_i and the marginal distributions F_i can be estimated by

$$p_i \simeq \hat{p}_i = \frac{m_i}{n}, \quad i = 1, \dots, d, \quad (4)$$

and

$$F_i(x) \simeq \hat{F}_i(x) = \frac{\sum_{j=1}^{m_i} 1(y_{i,j} \leq x)}{m_i}, \quad i = 1, \dots, d. \quad (5)$$

Similarly, the mean μ_i and variance σ_i^2 of F_i can be estimated from

$$\begin{aligned} \mu_i &\simeq \hat{\mu}_i = \frac{1}{m_i} \sum_{j=1}^{m_i} y_{i,j} \\ \sigma_i^2 &\simeq \hat{\sigma}_i^2 = \frac{1}{m_i} \sum_{j=1}^{m_i} (y_{i,j} - \hat{\mu}_i)^2. \end{aligned} \quad (6)$$

The estimation of the correlation matrix $\mathbf{r} = \{r_{ij}\}$, $i, j = 1, \dots, d$, corresponding to non-zero values of \mathbf{X} poses some difficulties since different coordinates of \mathbf{X} may be non-zero in different storms. Two options have been considered. First, select from the available record $(\mathbf{x}_1, \dots, \mathbf{x}_n)$ only those storms in which all coordinates are non-zero. This option is not viable since data shows that the resulting sample can be so short that reliable estimates of \mathbf{r} are not possible. Second, select from the available record $(\mathbf{x}_1, \dots, \mathbf{x}_n)$ all storms in which the entries of a particular pair (i, j) of coordinates are not zero and estimate r_{ij} from this record. The advantage of this approach is that allows more reliable estimates of \mathbf{r} . A potential problem is that the resulting estimate $\hat{\mathbf{r}}$ of \mathbf{r} may not be positive definite. We present in the following section a procedure for handling this situation. Let $\hat{\zeta}$ be the estimate of the matrix of correlation coefficients of the non-zero values of $\{X_i\}$ obtained from $\hat{\mathbf{r}}$ and Eq. 6. Since the differences between the correlation matrices $\boldsymbol{\rho}$ of the Gaussian image \mathbf{G} of \mathbf{X}_T and ζ are not significant for positively correlated random variables ([1], Section 3.1.1), we approximate $\boldsymbol{\rho}$ by $\hat{\zeta}$.

3.2 Monte Carlo algorithm

Suppose we need to generate n independent samples of \mathbf{X} . The proposed algorithm uses samples of \mathbf{X}_T as a substitute for samples of \mathbf{X} , and involves the following two steps.

Step 1. Generate n independent samples $(\mathbf{g}_1, \dots, \mathbf{g}_n)$ of \mathbf{G} with mean $\mathbf{0}$ and covariance matrix $\hat{\zeta}$.

Step 2. Calculate samples $(\mathbf{x}_{T,1}, \dots, \mathbf{x}_{T,n})$ of \mathbf{X}_T from $(\mathbf{g}_1, \dots, \mathbf{g}_n)$ and Eq. 3, and plot the resulting samples. It is assumed that all F_i are reverse Weibull distributions.

As previously stated, the generation of samples of \mathbf{G} may pose some difficulties since the estimate $\hat{\mathbf{r}}$ of the correlation matrix \mathbf{r} , and consequently the estimate $\hat{\zeta}$ of ζ , may not

be positive definite. The generation algorithm is based on the approximate representation

$$\mathbf{G} \simeq \tilde{\mathbf{G}} = \sum_{k=1}^{16} \nu_k^* V_k \phi_k \quad (7)$$

of \mathbf{G} , where $\{V_k\}$ are independent Gaussian variables with mean 0 and variance 1, $\{\nu_k, \phi_k\}$ denote the eigenvalues and the eigenvectors of $\hat{\boldsymbol{\zeta}}$, and $\nu_k^* = \nu_k$ if $\nu_k > 0$ and $\nu_k^* = 0$ otherwise. We use the approximation in Eq. 7 to generate samples of \mathbf{G} .

4 MATLAB functions

Two MATLAB functions have been developed,

`hurricane_dir_est.m` and
`hurricane_dir_mc.m`.

The first function estimates the parameters of the probability law of \mathbf{X}_T . The second function generate samples of \mathbf{X}_T . The dimension of \mathbf{X} is $d = 16$.

4.1 MATLAB function `hurricane_dir_est.m`

The input consists of:

- (1) A record at a specified milepost (see lines 23 to 27),
- (2) A range [cmin, cmax] of Weibull tail parameter c and the number nc of intervals in [cmin, cmax]. We note that cmax needs to be selected to avoid unrealistic tail parameters. It is suggested to set cmax = 10, and
- (3) A minimum number ncorr of non-zero pairs of non-zero readings needed to estimate entries of $\boldsymbol{\zeta}$. If ncorr is not reached for a pair (i, j) , we set $\hat{\zeta}_{ij} = 0$. It is suggested to set ncorr = 10.

The output consists of:

- (1) Estimates of the probabilities $p(i) = P(X_i = 0)$, $i = 1, \dots, d$,
- (2) Estimates of reverse Weibull parameters alpha1(i), c(i), and xi(i), $i = 1, \dots, d$,
- (3) Estimates zeta1(i, j) of the correlation coefficients ζ_{ij} , $i, j = 1, \dots, d$, and
- (4) Plots with estimates of the probabilities p_i ; mean, standard deviation, skewness of non-zero values of X_i ; estimates of the correlation coefficients of all data and of non-zero data; estimates of the parameters of the reverse Weibull distributions; and histograms of non-zero readings in all directions including Weibull densities fitted to these data.

The above output needs to be saved in a file for use in `hurricane_dir_mc.m`. The command `save estimates350 p zeta1 alpha1 c xi` may be used to store parameters needed for simulation. It is suggested that the file name be related to milepost number, for example, `estimates350` if dealing with milepost350.

4.2 MATLAB function `hurricane_dir_emc.m`

The input consists of:

- (1) A file with estimates of the parameters needed to define the probability law of \mathbf{X}_T , for example, the file `estimates350` and
- (2) The sample size `ns` and a seed `nseed` for sample generation.

The output consists of:

- (1) Three dimensional plots of the generated samples of \mathbf{G} and
- (2) Three dimensional plots and contour lines of the generated samples of \mathbf{X}_T .

5 Conclusions

A non-Gaussian model has been developed for hurricane wind speeds recorded in 16 equally spaced directions based on the theory of translation variables. A method has been presented for calibrating the wind model to site records. The calibrated model has been used to generate synthetic hurricane wind speeds of arbitrary length at a selected site.

References

- [1] M. Grigoriu. *Applied Non-Gaussian Processes: Examples, Theory, Simulation, Linear Random Vibration, and MATLAB Solutions*. Prentice Hall, Englewoods Cliffs, NJ, 1995.

Appendix A. MATLAB function hurricane_dir_est.m

```
function [p,mu,sig,gam3,zeta_t,zeta1,alpha1,c,xi] = ...
    hurricane_dir_est(cmin,cmax,nc,ncorr)
%
% It estimates:
%   (1) The probability  $p(i)=P(X_i=0)$  that coordinate
%        $i=1,\dots,16$  of wind speed is 0
%   (2) The mean  $\mu(i)$ , standard deviation  $\text{sig}(i)$ , and
%       skewness coefficient  $\text{gam3}(i)$  of the non-zero
%       values for each  $i=1,\dots,16$ 
%   (3) The correlation coefficients  $\{\text{zeta}_t(i,j)\}$ ,
%        $i,j=1,\dots,16$ , of the complete record,
%       i.e., including zero readings, and
%        $\{\text{zeta}_1(i,j)\}$ ,  $i,j=1,\dots,16$ , of
%       non-zero readings
%-----
% INPUT: (1) A record at a specified milepost
%         (see lines 23 to 27)
%         (2) Range [cmin,cmax] of Weibull tail
%             parameter c and nc = # of intervals
%             in [cmin,cmax]
%         NOTE: cmax is also used to limit the value
%               of the tail parameter, eg, cmax=10
%         (3) ncorr = the minimum number of non-zero
%             readings for which correlation is calculated
%             If ncorr is not reached, the correlation
%             coefficient is set 0
%             (Suggestion: Set ncorr=10)
%-----
% OUTPUT: (1) Estimates of  $\{p(i)\}$ ,  $i=1,\dots,16$ 
%          (2) Estimates of reverse Weibull parameters
%               $\{\text{alpha}_1(i), c(i), \text{xi}(i)\}$ ,  $i=1,\dots,16$ 
%          (3) Estimates of the correlation coefficients
%               $\{\text{zeta}_1(i,j)\}$ ,  $i,j=1,\dots,16$ , corresponding
%              non-zero wind speeds
%=====
% Load record = a (999,17)-matrix for a Milepost
% NOTE: THE FOLLOWING INSTRUCTION HAS TO BE MODIFIED
%       TO SELECT A DIFFERENT MILEPOST #
%-----
load milepost350;
q=matrix;
nr=length(q(:,1));
nu=mean_rate; % nu = the average number of hurricane/year
%              also in hppt://www.nist.gov/wind
%=====
% Estimates of probabilities  $p(i)$ 
% NOTE: All readings are  $\geq 0$ 
%-----
for i=1:16,
    p(i)=sum(q(:,i)>0)/nr;
end,
figure
plot(1:16,p)
```

```

xlabel('Wind direction')
ylabel('Estimates of probabilities of non-zero values')
%-----
% Construct non-zero wind speed records in each
% direction, estimate {mu(i), sig(i), gam3(i)}, and
% calculate coefficients of variation vq(i)=sig(i)/mu(i)
%-----
for i=1:16,
    nnz=0;
    for kr=1:nr,
        if q(kr,i)>0,
            nnz=nnz+1;
            xnz(nnz)=q(kr,i);
        end,
    end,
    xnzz=xnz(1:nnz);
    mu(i)=mean(xnzz);
    sig(i)=std(xnzz);
    vq(i)=sig(i)/mu(i);
    gam3(i)=mean(((xnzz-mu(i))/sig(i)).^3);
end,
figure
plot(1:16,mu,1:16,sig,':')
xlabel('Wind direction')
ylabel('Estimates of mean/std (solid/dotted lines) for non-zero values')
figure
plot(1:16,gam3)
xlabel('Wind direction')
ylabel('Estimates of skewness for non-zero values')
%-----
% Estimates of correlation coefficients
% {zeta_t(i,j)}, i,j=1,...,16
%-----
qq=q(:,1:16);
zeta_t=corrcoef(qq);
figure
mesh(1:16,1:16,zeta_t)
xlabel('Wind direction #')
ylabel('Wind direction #')
zlabel('Estimates of correlation coefficients \zeta_t')
%-----
% Estimates of correlation coefficients
% {zeta(i,j)}, i,j=1,...,16
%-----
for i=1:16,
    for j=1:16,
        q1=q(:,i);
        q2=q(:,j);
        nqq=0;
        for kr=1:nr,
            if q1(kr)>0 & q2(kr)>0,
                nqq=nqq+1;
                xqq(nqq,:)= [q1(kr) q2(kr)];
            end,
        end,
        if nqq<=01,
            zeta(i,j)=0;

```

```

        else,
            rr=corrcoef(xqq(1:nqq,1),xqq(1:nqq,2));
            rrr=rr(1,2);
            zeta(i,j)=rrr;
        end,
    end,
end,
figure
mesh(1:16,1:16,zeta)
xlabel('Wind direction #')
ylabel('Wind direction #')
zlabel('Estimates of correlation coefficients \zeta')
figure
contour(1:16,1:16,zeta)
xlabel('Wind direction #')
ylabel('Wind direction #')
title('Estimates of correlation coefficients \zeta')
%-----
%   Estimates of the paramters of reverse Weibull distributions
%   fitted to non-zero wind speeds (Method of moments)
%   USE [- RECORD] in all directions
%-----
%           Relationship between Weibull tail parameter
%           and skewness
%-----
dc=(cmax-cmin)/nc;
cc=cmin:dc:cmax;
lc=length(cc);
g1=gamma(1./cc+1);
g2=gamma(2./cc+1);
g3=gamma(3./cc+1);
skew=(g3-3*g1.*g2+2*g1.^3)./(g2-g1.^2).^(3/2);
% figure
% plot(cc,skew)
% xlabel('coefficient c')
% ylabel('skewness')
%-----
%           Calculation of skewness coefficients
%           for values of c>0 in [cmin,cmax]
%           and estimated tail parameters
%           {c(i)}, i=1,...,16
%-----
for i=1:16,
    muw(i)=-mu(i);
    sigw(i)=sig(i);
    gamw3(i)=-gam3(i);
    c(i)=interp1(skew,cc,gamw3(i),'spline');
%-----
%   NOTE: This condition is needed since
%           c can take very large values
%-----
    if c(i)>cmax,
        c(i)=cmax;
    end,
end,
end,
%-----
%           NOTE: If desired one or more or all c(i)'s

```

```

%                               can be assigned different values
%-----
for i=1:16,
    ggw1(i)=gamma(1./c(i)+1);
    ggw2(i)=gamma(2./c(i)+1);
    ggw3(i)=gamma(3./c(i)+1);
    alpha(i)=sigw(i)/sqrt(ggw2(i)-ggw1(i)^2);
    xi(i)=muw(i)-alpha(i)*ggw1(i);
end,
figure
plot(1:16,alpha,1:16,c,':',1:16,xi,'--')
xlabel('Wind direction #')
ylabel('Reverse Weibull parameters for non-zero readings')
title('Estimates of \alpha, c, and \xi (solid, dotted, and dashed lines)')
%-----
%   Plots of histograms and fitted reverse Weibull distributions
%   to non-zero wind speeds in all directions
%-----
for i=1:16,
    nnz=0;
    for kr=1:nr,
        if q(kr,i)>0,
            nnz=nnz+1;
            xnz(nnz)=q(kr,i);
        end,
    end,
    xnzz=xnz(1:nnz);
    figure
    hist_est(xnzz',1,30)
    hold
    yxi=xi(i):.1:50;
    yw=(yxi-xi(i))/alpha(i);
    fw=(c(i)/alpha(i))*(yw.^(c(i)-1)).*exp(-yw.^c(i));
    plot(-yxi,fw)
    xlabel('Wind speed (mph)')
    ylabel(['Direction ' int2str(i)])
    %   print
end,
zetal=zeta;
alpha1=alpha;
%=====
%   EXAMPLE:
%   [p,mu,sig,gam3,zeta_t,zetal,alpha1,c,xi]=hurricane_dir_est(.1,10,1000,10);
%   NOTE: Save the output needed for Monte Carlo simulation, e.g., use
%         save estimates350 p zeta alpha c xi
%         (estimates350 = file name, 350 since mileplot350 is used)

```

Appendix B. MATLAB function hurricane_dir_mc.m

```
function [thurr,xrw_mc,xrw_mc_ind,xrws_mc,xrws_mc_ind] = ...
hurricane_dir_mc(nyr,cws,nseed)
%
% INPUT FROM hurricane_dir_est.m ---> estimates1450_cw10 (for milepost1450),
% and consists of estimates of the parameters:
%
% * (alpha1, cw, xi) of reverse Weibull distributions
% fitted to non-zero wind speeds in 16 direction.
% * (alphas, xis) of reverse Weibull distributions
% fitted to non-zero wind speeds in 16 direction
% with imposed tail parameter cws = 10 (c = - 0.1)
% in all directions.
% * p = 16-dimensional vector with probabilities
% p(i)=P(X_i>0) of non-zero wind speeds.
% * zetal = (16,16) matrix of correlation coefficients
% for non-zero wind speeds.
%-----
% OTHER INPUT:
%
% * nyr = # of years required for simulation.
% * nseed = Monte Carlo simulation seed.
%-----
%
% OUTPUT:
%
% * thurr = times of thunderstorms in nyr years.
% * xrw_mc = generated wind speeds in 16 directions/nyr years
% using estimates of (alpha1, cw, xi), p(i), and zetal.
% * xrw_mc_ind = generated wind speeds in 16 directions/nyr years
% using estimates of (alpha1, cw, xi) and p(i) under the
% assumption that wind speeds in different directions
% are mutaully independent.
% * xrws_mc = generated wind speeds in 16 directions/nyr years
% using estimates of (alphas, xis), p(i), and zetal for
% an imposed tail parameter cws = - 1/c.
% * xrws_mc_ind = generated wind speeds in 16 directions/nyr years
% using estimates of (alphas, xis) and p(i) for an imposed
% tail parameter cws = - 1/c under the assumption that wind
% speeds in different directions are mutaully independent.
%=====
%
% REASONS FOR THE INDEPENDENCE ASSUMPTION AND THE RECOMMENDATION OF
% USING xrw_mc_ind; xrws_mc_ind RATHER THAN xrw_mc; xrws_mc
%
% (1) Correlation coefficients of all data (including 0's) are
% relatively small (maximum values are of order 0.7).
% (2) Correlation coefficients between random variables with
% finite probability mass at 0 provide limited information
% on the relationship between these random variables.
% (3) Estimates of the correlation coefficients of non-zero
% wind speeds can lead to inconsistencies, e.g., consider
% wind speed readings in 3 directions x(i,j), j=1,2,3,
% each of length n = 1000, and suppose the readings
% x(600:1000,1), x(1:400,2), x(800:1000,2), and x(1:600,3)
```

```

%           are zero. The estimates of the correlation coefficients
%           of these records are rho(1,2) not=0 (records x(:,1) & x(:,2)),
%           rho(2,3) not=0 (records x(:,2) & x(:,3)), but rho(1,3)=0
%           (records x(:,1) & x(:,3)).
%
%=====
% Load estimates delivered by hurricane_dir_est.m
% for a selected milepost (here milepost1450)
%-----
% load estimates350
load milepost1450
nu=mean_rate;
load estimates1450_cw10
nd=length(p);
%-----
% Total number of hurricanes in nyr years:
%     thurr = a vector with entries times at which
%             hurricanes occur in nyr years
%     nhurr = # of hurricanes in nyr years
%-----
rand('seed',nseed)
time=0;
ktime=0;
while time<=nyr,
    ktime=ktime+1;
    time=time-log(rand(1,1))/nu;
    thr(ktime)=time;
end,
nhurr=ktime-1;
thurr=thr(1:nhurr);
%-----
% Set 0 the entries of the matrices in which generated wind
% will be stores
%-----
xrw_mc=zeros(nhurr,16);
xrw_mc_ind=zeros(nhurr,16);
xrws_mc=zeros(nhurr,16);
xrws_mc_ind=zeros(nhurr,16);
%-----
% Generation of nhurr independent samples of a 16-dimensional
% standard Gaussian vector with covariance matrix zetal
%-----
% Construct an approximate spectral representation
% for a correlated standard Gaussian vector with
% covariance approximating zetal
%-----
[vzeta,dzeta]=eig(zetal);
ndd=0;
for kd=1:nd,
    if dzeta(kd,kd)>0,
        ndd=ndd+1;
        lamz(ndd)=dzeta(kd,kd);
        phiz(:,ndd)=vzeta(:,kd);
    end,
end,
%-----
% Generate required Gaussian samples

```

```

%-----
randn('seed',nseed);
gg=zeros(nhurr,nd);
for ks=1:nhurr,
    rg=randn(1,ndd);
    for kdd=1:ndd,
        gg(ks,:)=gg(ks,:)+lamz(kdd)*rg(kdd)*phiz(:,kdd)';
    end,
end,
gg=cdf('normal',gg,0,1);
%-----
% figure
% mesh(1:16,1:nhurr,gg)
% xlabel('Wind direction')
% ylabel('Sample number')
% zlabel('Gaussian image')
% xlim([1 16])
% set(gca,'xticklabel','')
% set(gca,'xtick',[1:16])
% set(gca,'xticklabel',[1:16])
% ylim([1 nhurr])
% set(gca,'yticklabel','')
% set(gca,'ytick',[1 10:10:nhurr])
% set(gca,'yticklabel',[1 10:10:nhurr])
% % print
%=====
% Translation from Gaussian to reverse Weibull space
% CASE 1: Estimates of (alpha, cw, xi), p(i), and zeta1
%-----
% gg=cdf('normal',gg,0,1);
for ks=1:nhurr,
    for i=1:nd,
        if gg(ks,i)>=1-p(i),
            uu=(gg(ks,i)-(1-p(i)))/p(i);
            xrw_mc(ks,i)=-xi(i)-icdf('wbl',uu,alpha1(i),cw(i));
        end,
% [ks i gg(ks,i) 1-p(i) xrw_mc(ks,i)]
% pause
    end,
end,
%*****
% UNDER INDEPENDENCE ASSUMPTION
%*****
for ks=1:nhurr,
    for i=1:nd,
        ur=rand(1,1);
        if ur>=1-p(i),
            uu=(ur-(1-p(i)))/p(i);
            xrw_mc_ind(ks,i)=-xi(i)-icdf('wbl',uu,alpha1(i),cw(i));
        end,
% [ks i gg(ks,i) 1-p(i) xrw_mc_ind(ks,i)]
% pause
    end,
end,
%-----
% Translation from Gaussian to reverse Weibull space
% CASE 2: Estimates of (alphas, xis), p(i), and zeta1

```

```

%           for an imposed tail parameter cws = - 1/c
%-----
%   gg=cdf('normal',gg,0,1);
for ks=1:nhurr,
    for i=1:nd,
        if gg(ks,i)>=1-p(i),
            uu=(gg(ks,i)-(1-p(i)))/p(i);
            xrws_mc(ks,i)=-xis(i)-icdf('wbl',uu,alphas(i),cws);
        end,
        [ks i gg(ks,i) 1-p(i) xrws_mc(ks,i)]
    %
    %   pause
    end,
end,
%*****
%   UNDER INDEPENDENCE ASSUMPTION
%*****
for ks=1:nhurr,
    for i=1:nd,
        ur=rand(1,1);
        if ur>=1-p(i),
            uu=(ur-(1-p(i)))/p(i);
            xrws_mc_ind(ks,i)=-xis(i)-icdf('wbl',uu,alphas(i),cws);
        end,
        [ks i gg(ks,i) 1-p(i) xrws_mc_ind(ks,i)]
    %
    %   pause
    end,
end,
%-----
% figure
% mesh(1:16,1:nhurr,xrw_mc)
% xlabel('Wind direction')
% ylabel('Sample number')
% zlabel('Simulated hurricane wind speeds')
% xlim([1 16])
% set(gca,'xticklabel','')
% set(gca,'xtick',[1:16])
% set(gca,'xticklabel',[1:16])
% ylim([1 nhurr])
% set(gca,'yticklabel','')
% set(gca,'ytick',[1 10:10:nhurr])
% set(gca,'yticklabel',[1 10:10:nhurr])
% %   print
%-----
% figure
% contour(1:16,1:ns,xrws_mc)
% xlabel('Wind direction')
% ylabel('Sample number')
% title('Simulated hurricane wind speeds')
% xlim([1 16])
% set(gca,'xticklabel','')
% set(gca,'xtick',[1:16])
% set(gca,'xticklabel',[1:16])
% ylim([1 nhurr])
% set(gca,'yticklabel','')
% set(gca,'ytick',[1 10:10:nhurr])
% set(gca,'yticklabel',[1 10:10:nhurr])
% %   print

```



```

% %-----
% figure
% contour(1:16,1:ns,xweib)
% xlabel('Wind direction')
% ylabel('Sample number')
% title('Simulated hurricane wind speeds')
% xlim([1 16])
% set(gca,'xticklabel','')
% set(gca,'xtick',[1:16])
% set(gca,'xticklabel',[1:16])
% ylim([1 ns])
% set(gca,'yticklabel','')
% set(gca,'ytick',[1 10:10:ns])
% set(gca,'yticklabel',[1 10:10:ns])
% % print
%=====
% [thurr,xrw_mc,xrw_mc_ind,xrws_mc,xrws_mc_ind]=hurricane_dir_mc(200000,10,123);

```