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**Designs for the Calibration
of Small Groups of Standards
in the Presence of Drift**

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DESIGNS FOR THE CALIBRATION OF SMALL GROUPS OF STANDARDS
IN THE PRESENCE OF DRIFT

by

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The process of calibrating a small number of "unknown" standards relative to one or two reference standards involved determining differences among the group of objects. Drift, due most often to temperature effects, or a "left-right" polarity effect can bias both the values assigned to the objects and the estimate of the effect of random errors. This note presents schedules of measurements of differences that eliminate the bias from these sources in the assigned value and variances and at the same time gives estimates of the magnitude of these extraneous components. The use of these designs in measurement process control is discussed and a computer program in BASIC is presented.

Key Words: Calibration; calibration design; experiment design; instrumental drift; measurement process; statistical analysis; trend elimination

1. Introduction

In very few processes can the effect of time be ignored. Instability in the object being measured, inability to maintain constant conditions or procedures, and variations in the detector or comparator all contribute to changes with time. A number of approaches have been suggested for reducing or eliminating the effects of these temporal effects on the validity of one's measurements. One way is to make measurements far enough apart in time (usually with some formal randomization procedure to guarantee statistical independence of the measurements) that the cumulative effects from the various sources appear in the random error component. At the other extreme, one can go to great lengths to eliminate these time dependent effects by achieving better environmental control, better instruments, better procedures, etc. If the measurements are to be transferred, as with instrument calibrations, then the first procedure leads to error bounds in which the random error limits include a between-time component whereas the latter procedure suppresses such a component. Neither of these represent the conditions of use adequately.

A compromise consists of arranging the experiment under its normal conditions so that it is as nearly as possible free of time dependent effects. The classic example of this is afforded by the calibration of thermometers in a bath with a gradually rising temperature using the schedule whose structure is as follows for a standard S, and 4 unknowns, T_1, T_2, T_3, T_4 .

S T_1 T_2 T_3 T_4 S T_4 T_3 T_2 T_1 S

If the measurements are evenly spread in time, then the average of the bath temperature for all thermometers are the same (see [3] for a discussion of this practice). A similar procedure has been followed in weighing where in the substitution method one measures in scale units

A, B, B+S, A+S

to obtain the difference A-B and the deflection corresponding to the sensitivity weight, S.

The calibration of a small number of "unknown" objects relative to one or two reference standards involves determining differences among the group of objects. Instrumental drift, due most often to temperature effects, or a "left-right" polarity effect can bias both the values assigned to the objects and the estimate of the effect of random errors. This note presents schedules of measurements of differences that eliminate the bias from these sources and at the same time gives estimates of the magnitude of these extraneous components. The use of these designs in measurement process control is discussed and a computer program in BASIC is presented in this report.

2. Measurement as a Process

A single isolated measurement, like a single event in history, is difficult to interpret unless it can be regarded as a part of a continuing process. When the measurement is looked upon as the output of a process--a production process whose output is the measured values--then one can attribute to the single measurement the properties of the process from which it arose (for a discussion of this approach, see Eisenhart [2]). Just as with any production process, the operating characteristics are determined by building some redundancy into the system. Redundancy is needed to assure oneself that he has indeed measured the sought after quantity, uncontaminated by extraneous factors related to the operator, instrument, environment, or other items.

Among the characteristics of the process are those associated with the ability to repeat a measurement both in the short term and in the long term. Repetitions made within a few hours, such as with designs having more observations than unknowns, usually exhibit less variation than those made at long time intervals. This additional long-term component of variance can be measured from the agreement among repeated measurements on the same quantity. In addition to these properties related to variability, one needs to incorporate checks on the systematic errors which may possibly affect the process, and, if possible, measurements that provide information as to the adequacy of the assumptions in the underlying physical model.

In calibration it is often convenient to measure a check standard along with the calibration of one or more unknowns. One thus has a value for monitoring the process that is on an equal footing with the unknowns. By tracking its long-run performance, one can determine not only the presence of components of variance, but also by recording ancillary information on environmental and other factors one can develop information for assessing the adequacy of the assumed physical model and for setting bounds to the effect from known sources of possible systematic error. This "check standard" need not be the value of a single item but may take the form of a difference between two such items or some linear combination of several.

The effect of some sources of systematic error can be eliminated by "balancing out" the effect by repeating the measurement of a difference, $(x - y)$ in the reverse order, $(y - x)$. Time dependent effects can be balanced out by using the techniques of this report. For others, it is sometimes possible to alter the conditions to levels of a factor beyond that known to have been in effect at the time of the measurement and to use the changes produced in the output at these extremes as a bounds to the effect of the factor.

In all cases one has to continually monitor the process output just as one does with an industrial production process if he is to have assurance that the calibrations are correct.

3. Substitution Weighing

Consider first the simple situation of scale deflections produced on a balance by adding weights A and B and a sensitivity weight S. One could use either of the following sequences.

<u>Sequence 1</u>	<u>Sequence 2</u>
A	A
A+S	B
B	B+S
B+S	A+S

In high precision work one invariably finds a change in balance response with time so that the value for the difference (A-B) will obviously be contaminated by whatever time effects exist for Sequence 1. If Sequence 2 is used, it may be represented as follows.

<u>Quantity</u>	<u>Effect of Drift</u>	<u>Scale Divisions</u>
A	-3Δ	x_1
B	$-\Delta$	x_2
B+S	Δ	x_3
A+S	3Δ	x_4

and the quantity

$$\frac{1}{2}(x_1 - x_2 - x_3 + x_4)$$

can be seen to give an unbiased value for (A-B) because the drift effect (a 2Δ change in scale reading between each observation) cancels out. The least squares values for A-B, S, and Δ in scale divisions are

$$\begin{aligned}(\hat{A-B}) &= \frac{1}{2}(x_1 - x_2 - x_3 + x_4) \\ \hat{S} &= \frac{1}{2}(x_1 - 3x_2 + 3x_3 - x_4) \\ \hat{\Delta} &= \frac{1}{4}(-x_1 + x_2 - x_3 + x_4)\end{aligned}$$

4. Thermometry

At NBS, the calibration of liquid in glass thermometers is usually carried out in a controlled bath which is continually heated so as to give a slight temperature increase with time. The temperature of the bath is measured by resistance thermometry at the start, middle, and end of the run with the test thermometers being run once each in the first interval and once again in reverse order in the second. The time sequence for the resistance measurements R_1, R_2, R_3 and the two series of test thermometer values denoted by $T_1 T_2 \dots T_k$ are as follows:

$$R_1 T_1 T_2 \dots T_k R_2 T'_k \dots T'_2 T'_1 R_3$$

If equal time intervals are maintained between readings of the test thermometers, then one would expect an increase, ΔT in temperature with each interval except perhaps the middle one in which the resistance thermometer reading, R_2 , is made. The analysis of this form of data is given in Appendix A.

5. Polarimeter Data

In determining the optical rotation of a quartz control plate used as reference standards in polarimeters, one measures the voltage response of a synchronous detector as the angle is varied. However, the response, y , of the system has a nearly linear drift with the angle so that one can represent this drift effect relative to the centroid of the data as being either $\dots -3\Delta, -2\Delta, -\Delta, 0, \Delta, 2\Delta, 3\Delta, \dots$ with Δ being the increment to the response added in each time interval. [For even n it is convenient to use $\dots -3\Delta, -\Delta, \Delta, 3\Delta \dots$ or 2Δ increment per time interval.]

In the polarimeter experiment the response is a linear function of angle so that the observation becomes

$$y_i = \alpha + \beta x_i + (i - \frac{n+1}{2})\Delta + \text{random error}$$

where the x_i are evenly spaced deviations from the nominal angle, e.g., $x = 0'', 10'', 20'', 30'', \dots$ If the usual estimate of α and β are to remain unbiased and unchanged in precision, then one must have

$$\sum x_i (i - \frac{n+1}{2}) = 0$$

so that the estimates are orthogonal to the drift in the detector. The following orderings have this property:

<u>Measurement Number</u>	<u>n = 4</u>		<u>n = 5</u>	
	<u>Quantity to be Measured</u>	<u>Setting for Polarimeter</u>	<u>Quantity to be Measured</u>	<u>Setting for Polarimeter</u>
1	$\alpha + 2\beta$	20"	$\alpha + \beta$	10"
2	α	0"	$\alpha + 4\beta$	40"
3	$\alpha + 3\beta$	30"	$\alpha + 2\beta$	20"
4	$\alpha + \beta$	10"	α	0"
5	---	---	$\alpha + 3\beta$	30"

6. Calibration Designs

The term calibration design has been applied [1] to experiments where only differences between nominally equal objects or groups of objects can be measured. Perhaps the simplest such experiment consists in measuring the differences between the two objects of the $n(n-1)/2$ distinct pairing that can be formed from n objects. Ordinarily the order in which these measurements are made is of no consequence. However, when the response of the comparator is time dependent, attention to the order is important if one wishes to minimize the effect of these changes. When this effect can be adequately represented by a linear drift, it is possible to balance out the effect by proper ordering of the observations. As with the polarimeter data, this drift can be represented by the series . . . -3Δ , -2Δ , $-\Delta$, 0 , Δ , 2Δ , 3Δ , . . . if $n(n-1)/2$ is odd or by . . . -5Δ , -3Δ , $-\Delta$, Δ , 3Δ , 5Δ . . . if $n(n-1)/2$ is even.

For $n = 4$, $n(n-1)/2 = 6$, and it turns out that it is not possible to balance out the drift effect with 6 measurements. However, with 8 measurements the balance can be achieved by the following order, denoting the four objects by A, B, C, D.

<u>Observation</u>	Observation is a Measurement of				
	<u>A</u>	<u>B</u>	<u>C</u>	<u>D</u>	<u>Δ</u>
y_1	+	-	0	0	-7
y_2	-	0	0	+	-5
y_3	0	0	+	-	-3
y_4	0	+	-	0	-1
y_5	0	+	0	-	1
y_6	-	0	0	+	3
y_7	+	0	-	0	5
y_8	0	-	+	0	7

The notation used here, the plus and minus signs, indicate the items entering into the difference measurement. Thus, y_2 is a measurement of the difference (D-A).

To see how the drift effect is balanced out, consider item C which occurs in the third, fourth, seventh, and eighth observations. In the third and eighth observations the item occurs positively and the corresponding drift effects are -3Δ and 7Δ respectively. For the fourth and seventh observations, item C occurs negatively while the corresponding

drift effects are $-\Delta$ and 5Δ . The overall effect can be represented by the sum of cross products of the columns for C and Δ , namely

$$[1](-3\Delta) + [-1](-\Delta) + [-1](5\Delta) + [1](7\Delta) = 0$$

using square brackets to denote the coefficient attached to the direction of the difference and parenthesis for the drift effect. For A, one has

$$[1](-7\Delta) + [-1](-5\Delta) + [-1](3\Delta) + [1](5\Delta) = 0$$

In general, if the cross products sum out to zero, then the drift effect is said to be completely "balanced out" or "orthogonal" to the items being measured.

7. Restraints

In calibration designs only differences between items are measured so that unless one or more of them are standards for which values are known, one cannot assign values for the remaining "unknown" items. Algebraically, one has a system of equations that is not of full rank and needs the value for one item or the sum of several items as the restraint to lead to a unique solution.

In the design of Section 6, for example, if one has a single standard and three unknowns, the standard can be assigned to any one of the letters. (The same would be true of three standards and one unknown.) If there are two standards and two unknowns, the choice of which pair of letters to assign for the standards is important in terms of minimizing the uncertainty in the unknown.

It turns out that the pairing of A with D or of B with C is slightly less efficient (see Appendix B) than the other pairings A with B or C with D. This results from the fact that the observation on the differences (A-D) and (B-C) are repeated and it is usually better (to achieve smaller variance for the test items) to measure differences between standards and unknowns than between pairs of standards.

8. Use of Calibration Design in Gage Block Calibration

The calibration design of Section 6 is used in gage block calibration at the National Bureau of Standards and the analysis and interpretation of the design for this application is representative of the principles involved in the use of the design in other applications.

At NBS two master sets of gage blocks are maintained for transferring length calibration to users gage blocks, these are designated A and B and their sum is designated by K. These are combined with two sets of unknowns, designated C and D. The difference (A-B) is used as the check standard.

If we denote the values determined for A B C and D by \hat{A} \hat{B} \hat{C} \hat{D} in accordance with the statisticians' practice of distinguishing the value from the experiment from the sought-after or long-run value, we may then write

$$\hat{A} = \frac{1}{24}(5y_1 - 2y_2 - y_3 - 2y_4 - 3y_5 - 2y_6 + 3y_7 + 2y_8) + \frac{K}{2}$$

$$\hat{B} = \frac{1}{24}(-5y_1 + 2y_2 + y_3 + 2y_4 + 3y_5 + 2y_6 - 3y_7 - 2y_8) + \frac{K}{2}$$

$$\hat{A-B} = \frac{1}{24}(10y_1 - 4y_2 - 2y_3 - 4y_4 - 6y_5 - 4y_6 + 6y_7 + 4y_8)$$

$$\hat{C} = \frac{1}{24}(-y_1 + 2y_2 + 5y_3 - 6y_4 - y_5 + 2y_6 - 7y_7 + 6y_8) + \frac{K}{2}$$

$$\hat{D} = \frac{1}{24}(y_1 + 6y_2 - 5y_3 - 2y_4 - 7y_5 + 6y_6 - y_7 + 2y_8) + \frac{K}{2}$$

where $\hat{A} + \hat{B}$ necessarily sum to K.

These values have the following standard deviations in terms of the long run precision as represented by the process standard deviation σ .

$$\text{s.d. } (\hat{A}) = \text{s.d. } (\hat{B}) = \sigma \sqrt{\frac{5}{48}}$$

$$\text{s.d. } (\hat{A-B}) = \sigma \sqrt{\frac{5}{12}}$$

$$\text{s.d. } (\hat{C}) = \text{s.d. } (\hat{D}) = \sigma \sqrt{\frac{13}{48}}$$

One also obtains values $\hat{\Delta}$ for Δ where

$$\hat{\Delta} = \frac{1}{168} (-7y_1 - 5y_2 - 3y_3 - y_4 + y_5 + 3y_6 + 5y_7 + 7y_8)$$

$$\text{s.d. } (\hat{\Delta}) = \sigma \sqrt{\frac{1}{168}}$$

Because this is an overdetermined system (more observations than unknowns) the deviation between observed and computed value is, in general, different from zero and reflects the random errors of measurement. These deviations, $d_1 d_2 \dots d_8$ are as follows:

$$d_1 = \frac{1}{168} (49y_1 - 7y_2 - 7y_3 + 21y_4 + 49y_5 + 49y_6 - 7y_7 + 21y_8)$$

$$d_2 = \frac{1}{168} (-7y_1 + 87y_2 + 13y_3 - 5y_4 + 33y_5 - 41y_6 + 53y_7 + 35y_8)$$

$$d_3 = \frac{1}{168} (-7y_1 + 13y_2 + 89y_3 + 25y_4 - 39y_5 + 37y_6 + 57y_7 - 7y_8)$$

$$d_4 = \frac{1}{168} (21y_1 - 5y_2 + 25y_3 + 111y_4 - 27y_5 + 3y_6 - 23y_7 + 63y_8)$$

$$d_5 = \frac{1}{168} (49y_1 + 33y_2 - 39y_3 - 27y_4 + 97y_5 + 25y_6 + 9y_7 + 21y_8)$$

$$d_6 = \frac{1}{168} (49y_1 - 41y_2 + 37y_3 + 3y_4 + 25y_5 + 103y_6 + 13y_7 - 21y_8)$$

$$d_7 = \frac{1}{168} (-7y_1 + 53y_2 + 57y_3 - 23y_4 + 9y_5 + 13y_6 + 73y_7 - 7y_8)$$

$$d_8 = \frac{1}{168} (21y_1 + 35y_2 - 7y_3 + 63y_4 + 21y_5 - 21y_6 - 7y_7 + 63y_8)$$

These deviations provide the information needed to obtain a value, s , which is the experiment's value for the process standard deviation, σ .

$$s = \sqrt{\frac{\sum (\text{dev})^2}{4}} \quad \text{degrees of freedom} = 4$$

The number of degrees of freedom results from taking the number of observations less the number of unknowns then adding one (for the restraint) to give $8 - 5 + 1 = 4$.

9. Example

Routine calibration of gage blocks is carried out with two NBS master blocks (designated S. and S..) and two test blocks (designated X and Y). The blocks are placed close together on a metal platen for a sufficiently long time to insure temperature equilibrium. A mechanical intercomparator is used to determine the difference between the blocks by first determining a reading for the block indicated by "+" then following with the block indicated by "-". The difference between these two readings is the observation, y (all values are in micro-inch). For a set of 0.101 in. blocks, the following data was obtained.

DATA FROM NBS CALIBRATION OF FOUR 0.101 INCH GAGE BLOCKS

<u>i</u>	<u>Schedule</u>	<u>Difference Measured</u>	<u>First Reading</u>	<u>Second Reading</u>	<u>Difference y(i)</u>	<u>Deviation</u>
1	+ - 0 0	S.-S..	52.0	52.5	-0.5	0.029
2	- 0 0 +	Y-S.	45.2	52.1	-6.9	-0.046
3	0 0 + -	X-Y	50.0	45.1	4.9	0.113
4	0 + - 0	S..-X	53.1	50.0	3.1	0.571
5	0 + 0 -	S..-Y	52.3	45.2	7.1	-0.238
6	- 0 0 +	Y-S.	45.1	52.0	-6.9	-0.079
7	+ 0 - 0	S.-X	52.0	50.1	1.9	-0.154
8	0 - + 0	S..-X	50.1	52.3	-2.2	0.304

S.+S.. = 6.4 used as restraint

S.-S.. = -0.133 used as check standard

$\sigma = .32$ accepted standard deviation

The values for the blocks and the drift effect, Δ , are

$$\hat{S}_. = \frac{1}{24}[5(-0.5) - 2(-6.9) - (4.9) - 2(3.1) - 3(7.1) - 2(-6.9) + 3(1.9) + 2(-2.2)] + \frac{(6.4)}{2}$$

$$= \frac{1}{24}(-6.0) + 3.2 = 2.9500$$

$$\hat{S}_{..} = \frac{1}{24}[-5(-0.5) + 2(-6.9) + (4.9) + 2(3.1) + 3(7.1) + 2(-6.9) - 3(1.9) - 2(-2.2)] + \frac{(6.4)}{2}$$

$$= \frac{1}{24}(6.0) + 3.2 = 3.4500$$

$$\hat{S} - S_{..} = -0.5$$

$$\begin{aligned}\hat{\bar{X}} &= \frac{1}{24} [-(-0.5) + 2(-6.9) + 5(4.9) - 6(3.1) - (7.1) + 2(-6.9) - 7(1.9) + \\ &\quad 6(-2.2)] + \frac{(6.4)}{2} \\ &= \frac{1}{24} (-54.8) + 3.2 = 0.9167\end{aligned}$$

$$\begin{aligned}\hat{\bar{Y}} &= \frac{1}{24} [(-0.5) + 6(-6.9) - 5(4.9) - 2(3.1) - 7(7.1) + 6(-6.9) - (1.9) + \\ &\quad 2(-2.2)] + \frac{(6.4)}{2} \\ &= \frac{1}{24} (-170.0) + 3.2 = -3.8833\end{aligned}$$

$$\begin{aligned}\hat{\Delta} &= \frac{1}{168} [-7(-0.5) - 5(-6.9) - 3(-4.9) - 1(3.1) + (7.1) + 3(-6.9) + 5(1.9) + \\ &\quad 7(-2.27)] \\ &= \frac{1}{168} (.7) = 0.0042\end{aligned}$$

The accepted standard deviation for the process is 0.32 μ -in so that one can compare the observed standard deviation, s ,

$$s = \sqrt{\Sigma \text{dev}^2 / 4} = \sqrt{\frac{.5208}{4}} = 0.361$$

to the accepted value by computing

$$F = \left(\frac{s}{\sigma}\right)^2 = \frac{0.1302}{0.1024} = 1.27$$

Had the ratio $(s/\sigma)^2$ exceeded 3.32 (the critical value for the 1% probability level of the F distribution), then the measurements would be regarded as being "out of control" and would be repeated. The other check on process performance is provided by the check standard for which the difference between $(\hat{S} - S_{..})$ and its accepted value should be less than 3 times the standard deviation of $(\hat{S} - S_{..})$. See Section 10 for a discussion of this test.

The drift term, $\hat{\Delta}$, has a standard deviation of $\sigma/\sqrt{168}$ or 0.025. The statistical significance of $\hat{\Delta}$ can be judged by forming the ratio $\frac{\hat{\Delta}}{\sigma/\sqrt{168}}$.

If this ratio exceeds 3, then $\hat{\Delta}$ would be regarded as significant. However, because the design has eliminated the effect of drift on the

values of the blocks, one would not be concerned about a "significant" $\hat{\Delta}$ unless it was greatly in excess of previously encountered values.

The deviations are computed as shown in Section 8, for example, for the deviation corresponding to y_8 is given by

$$\begin{aligned}(\text{dev})_8 &= \frac{1}{168} [21(-0.5) + 35(-6.9) - 7(4.9) + 63(3.1) + 21(7.1) - 21(-6.9) - \\ &\quad 7(1.9) + 63(-2.2)] \\ &= \frac{1}{168} [51.1] = 0.304\end{aligned}$$

10. Process Control

As mentioned in Section 2, continued monitoring of the measurement process is required to assure that predictions based on the accepted values for process parameters are still valid. For gage block calibration, the process is monitored for precision by comparison of the observed standard deviation to the accepted value, σ_w , by means of the F-test. In the case of the design of Section 6, the square of the ratio of the two standard deviations is compared to the critical value, $F(4, \infty, \alpha)$, which is the α probability point of the F distribution for degrees of freedom 4 and ∞ . [For calibrations at NBS, α is chosen as .01 to give $F(4, \infty, .01) = 3.32$].

The check for systematic error is given by comparison of the observed value of the difference, $S - S..$, between the two standards with its accepted value. The uncertainty of this difference is given by $\sigma_T = \sqrt{(5/12)\sigma_w^2 + \sigma_B^2}$ where σ_w is the "within run" standard deviation and σ_B is the component of variance arising from variations from run-to-run. The value of σ_T is obtained directly from the sequence of values of $S - S..$ arising in regular calibrations. The check standard test is therefore,

$$t = \frac{|\text{observed } (S - S..) - \text{accepted } (S - S..)|}{\sigma_T} < 3$$

i.e., t is compared to the critical value 3.0 which would correspond to the .003 probability level for the normal distribution.

If both the "precision" (F-test) and "accuracy" (t-test) criteria are satisfied, the process is regarded as being "in control" and values for the unknowns, X and Y , and their associated uncertainties are regarded as valid. Failure on either criterion is an "out-of-control" signal and the measurements are repeated.

When the between run component, σ_B , is present, the standard deviation associated with the values for the unknowns are given by

$$\begin{aligned}\sigma(S.) &= \sigma(S..) = \sqrt{\frac{5}{48}\sigma_w^2 + \sigma_B^2} = \sqrt{\sigma_T^2 - \frac{15}{48}\sigma_w^2} \\ \sigma(X) &= \sigma(Y) = \sqrt{\frac{13}{48}\sigma_w^2 + \sigma_B^2} = \sqrt{\sigma_T^2 - \frac{7}{48}\sigma_w^2}\end{aligned}$$

The value for the drift serves as an indicator of possible trouble if it changes markedly from its usual range of values. However, because any linear drift is balanced out, a change in the value does not of itself vitiate the results.

If the uncertainty attached to the restraint value is not negligible, this will lead to a possible systematic error in all measurements based on this restraint. Therefore, as a bound to this error one should, for the design of section 6, add to the uncertainty from random error an allowance of one-half the uncertainty in the sum ($S. + S..$). This is shown in the computer example.

11. Computer Program

Appendix C lists a computer program in BASIC for carrying out the calculation for the gage block example. The program can be used with any design provided one has the arrays of coefficients for the determination of the values of the unknowns and the deviations corresponding to the two arrays given in Section 8 for the gage block example.

The program calls for input of:

- a) Administrative data--designation of blocks, operator, date, etc.
- b) Process parameters--standard deviations, value for check standard, etc.
- c) Comparator readings

The computer programs provide in the output:

- a) Deviations, s.d.
- b) Values for unknowns, drift, and associated uncertainties.
- c) Statistical tests as to whether process can be regarded as "in control": on standard deviation and on value of check standard.

12. Other Designs for Elimination of Drift With Order of Measurement

The number of observations over which a linear drift could be expected to be valid varies with the type of measurement, but experience indicates that it is unusual if it is as large as 20. For all distinct pairings of n items $n(n-1)/2$ exceeds 20 for $n \geq 7$. The table below gives designs for $n = 5, 6, 7$ which are balanced for linear drift.

<u>n = 5</u> <u>10 Observations</u>	<u>n = 6</u> <u>18 Observations</u>	<u>n = 7</u> <u>21 Observations</u>
+ - 0 0 0 1-2	+ - 0 0 0 0 1-2	+ - 0 0 0 0 0 1-2
0 + - 0 0 2-3	0 + - 0 0 0 0 2-3	0 + - 0 0 0 0 0 2-3
0 0 + - 0 3-4	0 0 + - 0 0 0 3-4	0 0 + - 0 0 0 0 3-4
0 0 0 + - 4-5	0 0 0 + 0 - 4-6	0 0 0 + - 0 0 0 4-5
- 0 0 0 + 5-1	0 0 0 0 - + 6-5	0 0 0 0 + - 0 5-6
	- 0 0 0 + 0 5-1	0 0 0 0 0 + - 6-7
- 0 0 + 0 4-1		- 0 0 0 0 0 0 + 7-1
0 + 0 - 0 2-4	0 0 + 0 0 - 3-6	
0 - 0 0 + 5-2	- 0 0 + 0 0 4-1	0 + 0 - 0 0 0 0 2-4
0 0 + 0 - 3-5	0 + 0 0 - 0 2-5	0 0 + 0 - 0 0 0 3-5
+ 0 - 0 0 1-3	0 0 - 0 0 + 6-3	0 0 0 + 0 - 0 4-6
	+ 0 0 - 0 0 1-4	0 0 0 0 + 0 - 5-7
	0 - 0 0 + 0 5-2	- 0 0 0 0 + 0 6-1
		0 - 0 0 0 0 0 + 7-2
	0 0 0 - + 0 5-4	+ 0 - 0 0 0 0 0 1-3
	0 0 - 0 0 + 6-3	
	0 - 0 + 0 0 4-2	0 0 - 0 0 0 0 + 7-3
	- 0 + 0 0 0 3-1	0 - 0 0 0 + 0 6-2
	0 + 0 0 - 0 2-5	- 0 0 0 + 0 0 5-1
	+ 0 0 0 0 - 1-6	0 0 0 + 0 0 - 4-7
		0 0 + 0 0 - 0 3-6
		0 + 0 0 - 0 0 2-5
		+ 0 0 - 0 0 0 1-4

An alternate form of displaying the design is shown for $n = 5$ and is used for the other designs.

Designs that involve a subset of all possible pairings are given below:

<u>n = 6</u> <u>12 Observ.</u>	<u>n = 7</u> <u>14 Observ.</u>	<u>n = 8</u> <u>16 Observ.</u>	<u>n = 9</u> <u>18 Observ.</u>
1-2	1-2	1-2	1-2
5-1	2-3	2-3	2-3
2-3	3-4	3-4	6-5
4-6	4-5	4-5	3-1
3-4	5-6	5-6	5-4
6-5	6-7	6-7	8-9
	7-1	7-8	4-7
2-4		8-1	9-6
4-5	3-1		7-8
6-2	5-3	4-1	
3-1	7-5	7-4	7-1
1-6	2-7	2-7	4-6
5-3	4-2	5-2	9-7
	6-4	8-5	1-4
	1-6	3-8	3-9
		6-3	5-8
		1-6	6-3
			2-5
			8-2

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1. Bose, R. C. & Cameron, J. M., The Bridge Tournament Problem and Calibration Designs for Comparing Pairs of Objects, NBS J. of Res. 69B (1965), 323-332.
2. Eisenhart, C., Realistic Evaluation of the Precision and Accuracy of Instrument Calibration Systems, NBS J. of Res. 67C (1963), 161-187.
3. Swindells, J. F., Calibration of Liquid-in-Glass Thermometers, NBS Monograph 90, GPO, 1965.
4. Zelen, M., Linear Estimation and Related Topics, Chapter 17 of Survey of Numerical Analysis edited by J. Todd, McGraw Hill, New York City (1962), 558-584.

Thermometer Calibration

Liquid in glass thermometers are calibrated at NBS in a controlled bath in which the temperature is increasing in a nearly linear fashion with time. The temperature of the bath is measured by platinum resistance thermometry at the beginning, middle, and end of a run with the test thermometers being read once in the first interval and again in reverse order in the second interval. The time sequence for the resistance measurements, R_1, R_2, R_3 and the two series of thermometer values denoted by T'_1 and T_i are as follows:

$$R_1 \ T'_1 \ T'_2 \ \cdot \cdot \cdot \ T'_k \ R_2 \ T_k \ \cdot \cdot \cdot \ T_2 \ T_1 \ R_3$$

with uniform time intervals between the thermometer readings. Figure shows a schematic of the situation with the increment to the bath temperature being Δ for each time period except for the middle reading involving resistance thermometry where a step of α in temperature is assumed.

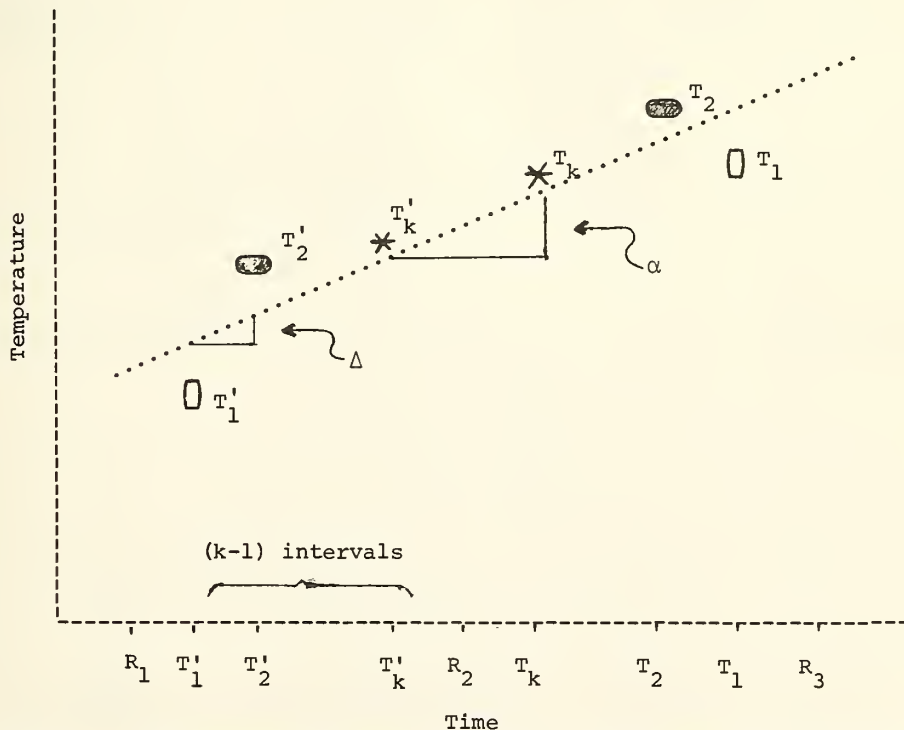


Figure Thermometer reading at fixed time intervals in a bath with linear drift.

The average $(T'_1 + T_1)/2$ will be the indication of the i -th thermometer at the temperature implied by $(R_1 + R_2 + R_3)/3$. The differences, $d_i = T_i - T'_1$ will be a measure of $\alpha + 2(k - i)\Delta$ so that the observational equations may be written

$$E(d) = E \begin{bmatrix} T_1 - T'_1 \\ T_2 - T'_2 \\ \cdot \\ \cdot \\ T_{k-1} - T'_{k-1} \\ T_k - T'_k \end{bmatrix} = \begin{bmatrix} \alpha + 2(k-1)\Delta \\ \alpha + 2(k-2)\Delta \\ \cdot \\ \cdot \\ \alpha + 2\Delta \\ \alpha \end{bmatrix} = \begin{bmatrix} 1 & 2(k-1) \\ 1 & 2(k-2) \\ \cdot & \cdot \\ \cdot & \cdot \\ 1 & 2 \\ 1 & 0 \end{bmatrix} \begin{bmatrix} \alpha \\ \Delta \end{bmatrix} = X \begin{bmatrix} \alpha \\ \Delta \end{bmatrix}$$

where X stands for the indicated matrix, and where $E(\)$ stands for the "expected value of," i.e., the limiting value if the effects of random error were eliminated.

The least squares estimates of α and Δ are given by the solution to the normal equations

$$(X'X) \begin{bmatrix} \alpha \\ \Delta \end{bmatrix} = X'd = \begin{bmatrix} \Sigma d \\ 2\Sigma d (k-i) \end{bmatrix}$$

where the inverse of the matrix of normal equations is

$$(X'X)^{-1} = \begin{bmatrix} k & k(k-1) \\ k(k-1) & 2k(k-1)(2k-1)/3 \end{bmatrix}^{-1} = \frac{1}{k(k^2-1)} \begin{bmatrix} 2(k-1)(2k-1) & -3(k-1) \\ -3(k-1) & 3 \end{bmatrix}$$

The estimates of α , Δ and σ^2 , the variance of the observations are given by

$$\hat{\alpha} = \frac{2}{k(k+1)} [3\Sigma id - (k+1)\Sigma d]$$

$$\hat{\Delta} = \frac{3}{k(k^2-1)} [(k+1)\Sigma d - 2\Sigma id]$$

$$\hat{\sigma}^2 = \frac{1}{k-2} [\Sigma d^2 - \hat{\alpha}\Sigma d - 2\hat{\Delta}\Sigma(k-i)d] = \frac{1}{k-2} \Sigma (\text{dev})^2$$

where $\text{dev}_i = d_i - \hat{\alpha} - 2(k-i)\hat{\Delta}$.

The standard deviation of the value for the test thermometer is $\sigma/\sqrt{2}$ and for α and Δ ,

$$s.d. (\alpha) = \sigma\sqrt{2(2k-1)/k(k+1)}$$

$$s.d. (\Delta) = \sigma\sqrt{3/k(k^2-1)}$$

Control on the measurement process is maintained by two forms of redundancy--one to check on the process average and the other to check on process variability. The first of these is provided by incorporating an NBS standard thermometer among the k thermometers and requiring that its value be within random error of its accepted value. The variability check is given by comparing $\hat{\sigma}$ with the long run value established for the process. When these conditions are satisfied, then one can regard the process as being in a state of control.

A typical set of data for this type of calibration is given in the following table. For simplicity the resistance measurements have been suppressed and the temperature reported directly.

Calibration of Thermometers
Data From NBS Calibration of 22 August 1972
Provided by J. Wise, NBS, Thermometry Section

<u>Thermometer</u>		<u>Observation</u>	<u>Averages</u>	<u>PRT - OBS =</u> <u>Correction at 40°</u>
Reference (PRT)		39.9378		
T' ₁		39.983	T ₁ 39.9870	-0.0436
T' ₂		39.913	T ₂ 39.9150	0.0284
T' ₃		39.966	T ₃ 39.9675	-0.0241
T' ₄	(check standard)	39.840	T ₄ 39.8410	0.1024*
Reference (PRT)		39.9422	PRT 39.9434	
T ₄		39.842		
T ₃		39.969		
T ₂		39.917		
T ₁		39.991	*accepted value is 0.1000	
Reference PRT		39.9501		

<u>i</u>	<u>d =</u> <u>T'_i - T_i</u>	<u>α</u>	<u>Δ</u>	<u>Predicted</u> <u>d</u>	<u>dev.</u>
1	-0.008	1	6	Σd = -0.077	-0.0071
2	-0.004	1	4	Σid = -0.033	-0.0052
3	-0.003	1	2	2/k(k+1) = 1/10	-0.0033
4	-0.002	1	0	3/k(k ² -1) = 1/20	-0.0014

$$\Sigma dev^2 = 0.00000270$$

$$\hat{\alpha} = \frac{1}{10} [3(-0.033) - 5(-0.017)] = -0.00140$$

$$\hat{\Delta} = \frac{1}{20} [5(-0.017) - 2(-0.033)] = -0.00095$$

$$\hat{\sigma} = \sqrt{\Sigma \text{dev}^2 / 2} = \sqrt{0.00000135} = 0.00115$$

$$\text{s.d. } (\hat{\alpha}) = \sigma \sqrt{\frac{7}{10}}$$

$$\text{s.d. (average T)} = \sigma / \sqrt{2}$$

$$\text{s.d. } (\hat{\Delta}) = \sigma \sqrt{\frac{1}{20}}$$

Least Squares Analysis of Calibration Designs

In this appendix the least squares analysis is presented in matrix form for those wishing to prepare a general analysis. Each formal statement will be illustrated by its application to the calibration design of Section 6.

It is assumed that the expected value of the observations represented in vector form as $y' = (y_1 \ y_2 \ . \ . \ . \ y_n)$ have expected values $E(y) = X\beta$ where β is the vector of parameters and X is the design matrix. It is also assumed that the errors of measurement are uncorrelated and have equal variance, i.e., $V(y) = \sigma^2 I$.

For the design of section 5,

$$X = \begin{bmatrix} 1 & -1 & 0 & 0 & -7 \\ -1 & 0 & 0 & 1 & -5 \\ 0 & 0 & 1 & -1 & -3 \\ 0 & 1 & -1 & 0 & -1 \\ 0 & 1 & 0 & -1 & 1 \\ -1 & 0 & 0 & 1 & 3 \\ 1 & 0 & -1 & 0 & 5 \\ 0 & -1 & 1 & 0 & 7 \end{bmatrix} \quad \beta = \begin{bmatrix} A \\ B \\ C \\ D \\ \Delta \end{bmatrix}$$

The matrix of normal equations is given by $(X'X)\beta = X'y$ which for calibration designs is not of full rank.

$$X'X\beta = \begin{bmatrix} 4 & -1 & -1 & -2 & 0 \\ -1 & 4 & -2 & -1 & 0 \\ -1 & -2 & 4 & -1 & 0 \\ -2 & -1 & -1 & 4 & 0 \\ 0 & 0 & 0 & 0 & 168 \end{bmatrix} \beta = X'y = \begin{bmatrix} 1 & -1 & 0 & 0 & 0 & -1 & 1 & 0 \\ -1 & 0 & 0 & 1 & 1 & 0 & 0 & -1 \\ 0 & 0 & 1 & -1 & 0 & 0 & -1 & 1 \\ 0 & 1 & -1 & 0 & -1 & 1 & 0 & 0 \\ -7 & -5 & -3 & -1 & 1 & 3 & 5 & 7 \end{bmatrix} y$$

In order to solve this system, a restraint in the form $h'\beta = K$ is imposed leading to the augmented equations (see Zelen [4]),

$$\begin{pmatrix} X'X & h \\ h' & 0 \end{pmatrix} \begin{pmatrix} B \\ \lambda \end{pmatrix} = \begin{pmatrix} X'y \\ K \end{pmatrix}$$

For the design as used in the calibration of gage blocks the restraint is that $A + B = K$, giving $h' = (1 \ 1 \ 0 \ 0 \ 0)$ and the augmented equations are

$$\begin{pmatrix} 4 & -1 & -1 & -2 & 0 & 1 \\ -1 & 4 & -2 & -1 & 0 & 1 \\ -1 & -2 & 4 & -1 & 0 & 0 \\ -2 & -1 & -1 & 4 & 0 & 0 \\ 0 & 0 & 0 & 0 & 168 & 0 \\ 1 & 1 & 0 & 0 & 0 & 0 \end{pmatrix} \beta = \begin{pmatrix} X'y \\ K \end{pmatrix}$$

The solution for the parameter values $\hat{\beta}$ are

$$\begin{pmatrix} \hat{\beta} \\ \lambda \end{pmatrix} = \begin{pmatrix} X'X & h \\ h' & 0 \end{pmatrix}^{-1} \begin{pmatrix} X'y \\ K \end{pmatrix} = \begin{pmatrix} C & g \\ g' & 0 \end{pmatrix} \begin{pmatrix} X'y \\ K \end{pmatrix}$$

where C is the indicated $K \times K$ matrix arising in the inversion process n . For the example

$$\begin{pmatrix} \hat{\beta} \\ \lambda \end{pmatrix} = \frac{1}{336} \begin{pmatrix} 35 & -35 & -7 & 7 & 0 & 168 \\ -35 & 35 & 7 & -7 & 0 & 168 \\ -7 & 7 & 91 & 21 & 0 & 168 \\ 7 & -7 & 21 & 91 & 0 & 168 \\ 0 & 0 & 0 & 0 & 2 & 0 \\ 168 & 168 & 168 & 168 & 0 & 0 \end{pmatrix} \begin{pmatrix} X'y \\ K \end{pmatrix} = \frac{1}{168} \begin{pmatrix} 35 & -14 & -7 & -14 & -21 & -14 & 21 & 14 & 84 \\ -35 & 14 & 7 & 14 & 21 & 14 & -21 & -14 & 84 \\ -7 & 14 & 35 & -42 & -7 & 14 & -49 & 42 & 84 \\ 7 & -14 & -35 & 42 & -7 & 14 & 49 & -42 & 84 \\ -7 & -5 & -3 & -1 & 1 & 3 & 5 & 7 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \end{pmatrix} \begin{pmatrix} y \\ K \end{pmatrix}$$

$$\text{for which } C = \frac{1}{336} \begin{pmatrix} 35 & -35 & -7 & 7 & 0 \\ -35 & 35 & 7 & -7 & 0 \\ -7 & 7 & 91 & 21 & 0 \\ 7 & -7 & 21 & 91 & 0 \\ 2 & 0 & 0 & 0 & 2 \end{pmatrix}$$

The variances of the parameters are given by $C_{ii}\sigma^2$ and of linear functions, $\ell'\beta$, the variance is $\ell'C\ell\sigma^2$. For the example

$$V(\hat{A}) = V(\hat{B}) = C_{11}\sigma^2 = C_{22}\sigma^2 = 35\sigma^2/336 = 5\sigma^2/48$$

$$V(\hat{C}) = V(\hat{D}) = C_{33}\sigma^2 = C_{44}\sigma^2 = 91\sigma^2/336 = 13\sigma^2/48$$

$$V(\hat{A}-\hat{B}) = (C_{11} + C_{22} - 2C_{12})\sigma^2 = 140\sigma^2/336 = 5\sigma^2/12$$

$$V(\hat{A}+\hat{B}) = 0$$

$$V(\hat{C}+\hat{D}) = (C_{33} + C_{44} + 2C_{34})\sigma^2 = 224\sigma^2/336 = 2\sigma^2/3$$

$$V(\hat{C}-\hat{D}) = (C_{33} + C_{44} - 2C_{34})\sigma^2 = 140\sigma^2/336 = 5\sigma^2/24$$

$$V(\hat{\Delta}) = C_{55}\sigma^2 = 2\sigma^2/336 = \sigma^2/168$$

If one had assigned the two standards to positions B and C instead of to A and B as was done, then one would be repeating the measurement of the difference (B-C), and of the difference (A-D). These differences are internal to the pair of standards over the pair of unknowns and one might suspect that they add little to the transfer from standard to test item. This is confirmed by examination of the inverse of the matrix of normal equations,

$$\begin{pmatrix} 4 & -1 & -1 & -2 & 0 & 0 \\ -1 & 4 & -2 & -1 & 0 & 1 \\ -1 & -2 & 4 & -1 & 0 & 1 \\ -2 & -1 & -1 & -14 & 0 & 0 \\ 0 & 0 & 0 & 0 & 168 & 0 \\ 0 & 1 & 1 & 0 & 0 & 0 \end{pmatrix}^{-1} = \frac{1}{168} \begin{pmatrix} 56 & 0 & 0 & 28 & 0 & 84 \\ 0 & 14 & -14 & 0 & 0 & 84 \\ 0 & -14 & 14 & 0 & 0 & 84 \\ 28 & 0 & 0 & 56 & 0 & 84 \\ 0 & 0 & 0 & 0 & 1 & 0 \\ 84 & 84 & 84 & 84 & 0 & 0 \end{pmatrix}$$

The variances for the standards are

$$v(\hat{B}) = v(\hat{C}) = 14\sigma^2/168 = \sigma^2/12$$

$$v(\hat{B}-\hat{C}) = 56\sigma^2/168 = \sigma^2/3$$

which is smaller than for the restraint $A + B = K$.

However, for the test items the variances are

$$v(\hat{A}) = v(\hat{D}) = 56\sigma^2/168 = \sigma^2/3$$

which is larger than that with the restraint $A + B = K$ for which the corresponding variance is $13\sigma^2/48$.

The estimate for the test item, A, is

$$\hat{A} = \frac{1}{24} [8y_1 - 4y_2 - 4y_3 - 4y_5 - 4y_6 + 8y_7] + \frac{K}{2}$$

which does not involve y_4 and y_8 which are measurements of the difference between the two standards, i.e., of B-C. Thus, there is a gain in efficiency in the calibration of the test block by using positions A and B for the standards, the efficiency factor being $(\sigma^2/3)/(13\sigma^2/48) = 16/13$.

FOOTNOTE 2

If there were but a single standard, A, the inverse of the matrix of normal equations would be

$$\begin{bmatrix} 4 & -1 & -1 & -2 & 0 & 1 \\ -1 & 4 & -2 & -1 & 0 & 0 \\ -1 & -2 & 4 & -1 & 0 & 0 \\ -2 & -1 & -1 & 4 & 0 & 0 \\ 0 & 0 & 0 & 0 & 168 & 0 \\ 1 & 0 & 0 & 0 & 0 & 0 \end{bmatrix}^{-1} = \frac{1}{168} \begin{bmatrix} 0 & 0 & 0 & 0 & 0 & 168 \\ 0 & 70 & 42 & 28 & 0 & 168 \\ 0 & 42 & 70 & 28 & 0 & 168 \\ 0 & 28 & 28 & 56 & 0 & 168 \\ 0 & 0 & 0 & 0 & 1 & 0 \\ 168 & 168 & 168 & 168 & 0 & 0 \end{bmatrix}$$

The variances of the test items are

$$V(\hat{B}) = V(\hat{C}) = 70\sigma^2/168 = 5\sigma^2/12$$

$$V(\hat{D}) = 56\sigma^2/168 = \sigma^2/3$$

FOOTNOTE 3

If the sum of all four were taken as the restraint, the inverse of the matrix of normal equations would be

$$\begin{bmatrix} 4 & -1 & -1 & -2 & 0 & 1 \\ -1 & 4 & -2 & -1 & 0 & 1 \\ -1 & -2 & 4 & -1 & 0 & 1 \\ -2 & -1 & -1 & 4 & 0 & 1 \\ 0 & 0 & 0 & 0 & 168 & 0 \\ 1 & 1 & 1 & 1 & 0 & 0 \end{bmatrix}^{-1} = \frac{1}{336} \begin{bmatrix} 49 & -21 & -21 & -7 & 0 & 84 \\ -21 & 49 & -7 & -21 & 0 & 84 \\ -21 & -7 & 49 & -21 & 0 & 84 \\ -7 & -21 & -21 & 49 & 0 & 84 \\ 0 & 0 & 0 & 0 & 2 & 0 \\ 84 & 84 & 84 & 84 & 0 & 0 \end{bmatrix}$$

The variances of all four test items are the same

$$V(\hat{A}) = V(\hat{B}) = V(\hat{C}) = V(\hat{D}) = 49\sigma^2/336 = 7\sigma^2/48$$

Computer Program for Analysis of Gage Block Data

```

5 *****G A G E      B L O C K      C A L I B R A T I O N*****
10 *   THIS PROGRAM COMPUTES THE VALUES OF THE UNKNOWN GAGE BLOCKS,
15 *   AND PERFORMS TWO STATISTICAL TESTS, THE F-TEST AND THE T-TEST,
20 *   TO DETERMINE IF THE PROCESS IS IN CONTROL.
25 *****
30 *                               INPUT
35 *
40 *   THIS PROGRAM CALLS FOR A USER CREATED DATA FILE WHICH CONSISTS
45 *   OF THE FOLLOWING:
50 *   (1) (C$(I),I=1,3) - DATE,OBSERVER,INSTRUMENT
55 *   (2) K - VALUE OF RESTRAINT(MICROINCHES)
60 *   (3) N0 - NOMINAL SIZE OF TEST BLOCKS(INCHES)
65 *   (4) S5 - ACCEPTED VALUE OF CHECK STANDARD(MICROINCHES)
70 *   (5) B5 - ACCEPTED S.D. OF THE INSTRUMENT(MICROINCHES)
72 *   (6) B6 - ACCEPTED TOTAL S.D.
75 *   (7) M2 - UNCERTAINTY IN THE RESTRAINT
80 *   (8) X(I),Y(I) - OBSERVED READINGS (EIGHT PAIRS)
85 *****
90 *
95 *
100 *****
105 *
110 *   DATA VALUES WHICH ARE DETERMINED BY THE CALIBRATION DESIGN,
115 *   AND WHICH ARE STORED WITHIN THIS PROGRAM IN DATA STATEMENTS,
120 *   ARE AS FOLLOWS:
125 *   (1) R(I,J) - LEAST SQUARES COEFF TO COMPUTE THE UNKNOWNNS
130 *   (2) M(I,J) - LEAST SQUARES COEFF TO COMPUTE THE DEVIATION
135 *   (3) E(I) - VARIANCE FACTOR
138 *   (4) F(I) - DRIFT VECTOR
140 *   (5) C1 - MATRIX DIVISOR
142 *   (6) G$(I) - BLOCK DESIGNATION
145 *
150 *   OTHER VARIABLES:
155 *   (1) N - NO. OF BLOCKS IN THE CALIBRATION (N = 4)
160 *   (2) G1 - NO. OF OBSERVATIONS (G1 = 8)
165 *   (3) F4 - F RATIO (CRITICAL VALUE FOR P = .01); (F4 = 3.32)
170 *   (4) A4 - NO. OF DEGREES OF FREEDOM (A4 = 4)
175 *****
180 *
185 *
190 *****
195 *                               OUTPUT
200 *
205 *   (1)   DATE,OBSERVER,INSTRUMENT
210 *   (2)   COMPARATOR READINGS
215 *   (3)   OBSERVED DIFFERENCES
220 *   (4)   DEVIATIONS
225 *   (5)   VALUES OF THE UNKNOWNNS
230 *   (6)   OBSERVED STANDARD DEVIATION
235 *   (7)   STATISTICAL TESTS
240 *   (8)   UNCERTAINTY STATEMENT
245 *****

```

```

250 DIM C$(3),B(4,9),M(8,8),A(9),X(8),Y(8),G$(4)
255 DATA 35,-14,-7,-14,-21,-14,21,14,84
260 DATA -35,14,7,14,21,14,-21,-14,84
265 DATA -7,14,35,-42,-7,14,-49,42,84
270 DATA 7,42,-35,-14,-49,42,-7,14,84
275 DATA 49,-7,-7,21,49,49,-7,21
280 DATA -7,87,13,-5,33,-41,53,35
285 DATA -7,13,89,25,-39,37,57,-7
290 DATA 21,-5,25,111,-27,3,-23,63
295 DATA 49,33,-39,-27,97,25,9,21
300 DATA 49,-41,37,3,25,103,13,-21
305 DATA -7,53,57,-23,9,13,73,-7
310 DATA 21,35,-7,63,21,-21,-7,63
312 DATA -7,-5,-3,-1,1,3,5,7
315 DATA .31250,.31250,.145833,.145833
320 DATA 168
322 DATA S.,S.,X,Y
330 N=4
335 G1=8
340 F4=3.32
345 A4=4
350 R1=G1*1
355 * READ COEFFICIENTS USED TO COMPUTE VALUES OF THE BLOCKS
360 FOR I = 1,N
365 FOR J1 = 1,R1
370 READ B(I,J1)
375 NEXT J1
380 NEXT I
385 * READ COEFFICIENTS USED TO COMPUTE THE DEVIATIONS
390 FOR I = 1,G1
395 FOR J1= 1,G1
400 READ M(I,J1)
405 NEXT J1
410 NEXT I
411 * READ DRIFT VECTOR
412 FOR I = 1,G1
413 READ F(I)
414 NEXT I
415 * READ VARIANCE VECTOR
420 FOR I = 1,N
425 READ E(I)
430 NEXT I
435 * READ MATRIX DIVISOR
440 READ C1
441 * READ BLOCK DESIGNATIONS
442 FOR I = 1,N
443 READ G$(I)
444 NEXT I

```

```

445 * DEFINE USER DATA FILE - DATA1
450 FILES DATA1
455 * READ ADMINISTRATIVE DATA AND PROCESS PARAMETERS
460 READ#1,CS(1),CS(2),CS(3)
465 READ#1,K,N0,S5,B5,B6,M2
470 * READ COMPARATOR READINGS AND COMPUTE THEIR DIFFERENCES
471 * ALSO, COMPUTE DRIFT =D1, AND S.D.(DRIFT) = S1
473 D1=0
475 FOR I = 1,G1
480 READ#1,X(I),Y(I)
485 A(I)=X(I)-Y(I)
487 D1=D1+A(I)*F(I)
490 NEXT I
492 D1=D1/C1
493 S1=B5*(1./C1)↑.5
495 * SET A(9)= RESTRAINT
500 A(9)=K
505 * COMPUTE VALUES = V(I), S.D. = Z(I), AND UNCERTAINTY = C(I)
510 FOR I = 1,N
515 Y1=0
520 FOR J1=1,R1
525 Y1=Y1+B(I,J1)*A(J1)
530 NEXT J1
535 V(I)=Y1/C1
540 Z(I)=(B6↑2-B5↑2*E(I))↑.5
545 C(I)=3*Z(I)*.5*M2
550 NEXT I
555 * COMPUTE CHECK STANDARD
560 C5=V(1)-V(2)
565 * COMPUTE THE DEVIATIONS AND THE OBSERVED S.D.
570 S0=0
575 FOR I = 1,G1
580 D0=0
585 FOR J1=1,G1
590 D0=D0+M(I,J1)*A(J1)
595 NEXT J1
600 D(I)=D0/C1
605 S0=S0+D(I)↑2
610 NEXT I
615 S=(S0/A4)↑.5
620 * PERFORM STATISTICAL TESTS
625 F=(S/B5)↑2
630 T=(C5-S5)/B6

```



```

635*****J U T P U T*****
640 PRINT,906
645 PRINT DATE ,CS(1)
650 PRINT OBS. ,CS(2)
655 PRINT INSTR.,CS(3)
660 PRINT,905
665 PRINT OBSERVATIONS
670 FOR I = 1,G1
675 PRINT X(I),Y(I)
680 NEXT I
685 *PRINT OBSERVED DIFFERENCES AND DEVIATIONS
690 PRINT,695
695 FMT //,X18,A(I),X8, DEV(I)
700 FOR I = 1,G1
705 PRINT,710,A(I),D(I)
710 FMT X14,F9.3,X4,F9.3
715 NEXT I
720 * PRINT VALUES OF THE BLOCKS
722 PRINT,723
723 FMT //,X58,UNCERTAINTY
725 PRINT,730
730 FMT X19,NOM.,X8,CORR.,X10,S. D.,X4,3(S.D.)*.5(S.E.)
735 FOR I = 1,N
740 PRINT,745,G(I),NO,V(I),Z(I),C(I)
745 FMT F12.6,F11.2,X7,F8.5,F15.5
750 NEXT I
755 * PRINT STATISTICAL INFORMATION
760 PRINT,765
765 PRINT OBS. S.D,ACC. S.D.,F TEST,F RATIO,D.F.
770 PRINT,775,S,B5,F,F4,A4
775 FMT F7.4,X5,F9.5,X3,F8.3,F12.2,I10
780 IF F>F4 THEN 790
785 GO TO 800
790 PRINT,795
795 PRINT *****S. D. IS NOT IN CONTROL*****
800 PRINT,805
805 PRINT OBS. CHECK,ACC. CHECK,T TEST
810 PRINT,815,C5,S5,T
815 FMT F10.5,F11.5,F12.5
820 IF ABS(T) > 3 THEN 830
825 GO TO 840
830 PRINT,835
835 PRINT *****CHECK STANDARD IS NOT IN CONTROL*****
840 PRINT,905
845* PRINT DRIFT AND S.D.(DRIFT)
885 PRINT,890,D1,S1
890 FMT DRIFT = ,F10.4/ S.D.(DRIFT) = F10.4
895 PRINT,900,K
900 FMT ///, RESTRAINT (S.*S..) = ,F8.2
901 PRINT,903,M2
902 PRINT,906
903 FMT SYSTEMATIC ERROR(S.E.) IN RESTRAINT = ,F8.2
905 FMT //
906 FMT /////
910 STOP
915 END

```

INPUT = USER DATA FILE

10 MAY 28 1974, HOWELL, FEDERAL 1
 20 6.4, .101, -.133, .32, .49, .20
 30 52.0, 52.5
 40 45.2, 52.1
 50 50.0, 45.1
 60 53.1, 50.0
 70 52.3, 45.2
 80 45.1, 52.0
 90 52.0, 50.1
 100 50.1, 52.3

OUTPUT

DATE MAY 28 1974
 OBS. HOWELL
 INSTR. FEDERAL 1

OBSERVATIONS

52	52.5
45.2	52.1
50	45.1
53.1	50
52.3	45.2
45.1	52
52	50.1
50.1	52.3

A(I)	DEV(I)
-.500	.029
-6.900	-.046
4.900	.113
3.100	.571
7.100	-.237
-6.900	-.079
1.900	-.154
-2.200	.304

	NOM.	CORR.	S. D.	UNCERTAINTY 3(S.D.) + .5(S.E.)
S.	.101000	2.95	.45618	1.46854
S..	.101000	3.45	.45618	1.46854
X	.101000	.92	.47452	1.52355
Y	.101000	-3.88	.47452	1.52355

OBS. S.D.	ACC. S.D.	F TEST	F RATIO	D.F.
.3607	.32000	1.271	3.32	4

OBS. CHECK	ACC. CHECK	T TEST
-.50000	-.13300	-.74898

DRIFT = .0042
 S.D.(DRIFT) = .0247

RESTRAINT (S.+S..) = 6.40
 SYSTEMATIC ERROR(S.E.) IN RESTRAINT = .20

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