## NATIONAL BUREAU OF STANDARDS REPORT

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Statistical Engineering Laboratory National Eureau of Standards Washington, D. C. 20234

TABLES USEFUL IN ESTIMATING THE MEAN VALUE FUNCTION OF A FUNDAMENTAL RANDOM PROCESS

by

I. R. Savage and E. Lukacs



**U. S. DEPARTMENT OF COMMERCE NATIONAL BUREAU OF STANDARDS**  **U. S. DEPARTMENT OF COMMERCE** 

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#### TABLES USEFUL IN ESTIMATING THE MEAN VALUE FUNCTION OF A FUNDAMENTAL RANDOM PROCESS\*

by

#### I. Richard Savage and Eugene Lukacs National Bureau of Standards

1. Introduction. Let y(t) be a stochastic process and assume that

- (1.1) y(t) = x(t) + f(t)
- (1.2)  $f(t) = K_1 \mathscr{O}_1(t) + K_2 \mathscr{O}_2(t) + \dots + K_s \mathscr{O}_s(t)$

where x(t) is a fundamental random process.

In a forthcoming paper [4] H. B. Mann assumes that the y(t) process is observed over an interval [0,T] and developes a method for finding the maximum likelihood estimates for the parameters  $K_1, K_2, \ldots, K_s$  of the mean value function  $f(t)^{**}$ . In a technical note [5] he gives a brief exposition of his results and also explicit formulae for the practically important case where f(t) is a polynomial in t. As in least squares estimates it seems also here advantageous to use orthogonal polynomials. However, it is here necessary to assume that the derivatives  $\emptyset'_{i}(t)$  are the first s polynomials of a system orthogonal with respect to the weight function 1 over an interval [0,T]. The functions  $\emptyset_{j}(t)$  becomethen the integrals of Legendre polynomials adapted to the interval [0,1]. As a consequence the mean value function must have initial and terminal value zero, i.e.,

$$(1.3) f(0) = f(T) = 0.$$

While (1.3) will apply in certain physical situations its failure will in general prevent the use of orthogonal polynomials in estimating a mean value function \*\*\*. It will then be necessary

\*The preparation of this paper was sponsored by the U.S. Naval Ordnance Test Station, Inyokern.

\*\*The  $\emptyset_j(t)$  are known functions, and are subject to certain restrictions, stated in [4] and [5].

\*\*\*Ordinarily the condition (1.3) could be obtained by a rotation of axis, but in the situation under consideration such a rotation would disturb the nature of the random process, and hence can not be used. to assume that  $\emptyset(t) = t^{j}$  and to use the methods developed for this case.

Let us therefore suppose that

(1.4) 
$$f(t) = K_1 t + K_2 t^2 + \dots + K_s t^s$$

According to the formulae given in [4] and [5] the estimate  $\hat{K}_{j}$  of  $K_{j}$  is given by

(1.5) 
$$\bigwedge_{K_{j}}^{\Lambda} = \sum_{r=1}^{s} \Phi^{jr} r \int_{0}^{T} t^{r-1} dy(t) \qquad (j = 1, 2, ..., s)$$

Here  $\oint^{jr} = \frac{s^{jr}}{jrT^{j+r-1}}$  where the matrix  $||S^{jr}||$  is the inverse of the matrix  $\left\|\frac{1}{j+r-1}\right\|_{j,r} = 1, \ldots, s$ . The purpose of the present note is to simplify the computations necessary for obtaining the estimates (1.5). This is done by studying first the matrix  $\left\|\frac{1}{1+j-1}\right\|$  and giving formulae for finding its inverse. Tables of the coefficients  $S_{n}^{jr}$  are then computed for n = 2(1)10. These tables will be found at the end of this report. The same matrix occurs in least square theory when an integral is minimized instead of a sum. The inversion of this and related matrices was studied by A. R. Collar [2], [3] it is however believed that the proofs given here are somewhat more elementary.

2. A theorem due to A. Cauchy. The following theorem is due to A. Cauchy [1]. Let  $a_1, \ldots, a_n$ ,  $b_1, \ldots, b_n$  be 2n numbers and consider the determinant whose elements are of the form  $1/(a_1+b_k)$ (i,k = 1,2,...,n). Then

(2.1) 
$$\left\|\frac{1}{a_{j}+b_{k}}\right\|_{i,k=1,\ldots,n} = \frac{\frac{1 \cdot \cdot n}{||}}{\frac{j>k}{j>k}} \frac{(a_{j}-a_{k})(b_{j}-b_{k})}{\frac{1 \cdot \cdot n}{||}}$$

This theorem as well as an indication of its proof may also be found in [7], page 98, problem 3.

3. <u>Inversion of a certain matrix</u>. In the introduction we proposed to consider the matrix (3.1)  $S_n = \left\| \frac{1}{i+j-1} \right\|_{i,j=1,...,n}$  and to find its inverse. Clearly the determinant of  $S_n$  as well as all its minors are of the form discussed in Section 2. We can therefore compute the inverse  $S_n^{-1}$  by applying Cauchy's theorem. We denote by  $\Delta_n^{ij}$  the minor of the element in the i-th row and and j-th column of the determinant  $\Delta_n$  of the matrix  $S_n$ . If we write  $S_n^{ij}$  for the element in the i-th row and j-th column of the inverse  $S_n^{-1}$  then

(3.2) 
$$S_{n}^{ij} = (-1)^{i+j} \frac{\Delta_{n}^{ij}}{\Delta_{n}}$$

If we use (2.1) to find  $\Delta_n^{jj}$  and  $\Delta_n$  we obtain by an elementary computation (n-1)

(3.3) 
$$\Delta_{n} = \frac{\left(\frac{1}{k!} + \frac{1}{k!}\right)^{-1}}{\frac{1}{n!} \frac{n-1}{(n^{2}-k^{2})^{n-k}}}$$

and

(3.4) 
$$\Delta_{n}^{ij} = \frac{(n+i-1)!(n+j-1)!}{[(i-1)!(j-1)!]^{2}n!(n-i)!(n-j)!} \frac{\binom{1}{k=1}}{\binom{n-1}{1}} \frac{-1}{j+i+1}$$

 $\left[\frac{n-1}{2}\right]_{3}$ 

It is easy to show inductively that

(3.5) 
$$\frac{n^n}{n!} \prod_{k=1}^{n-1} \frac{(n^2 - k^2)^{n-k} k!}{(n+k)!} = 1$$

so that we obtain from (3.2), (3.3) and (3.4)

(3.6) 
$$S_n^{ij} = \frac{(-1)^{i+j}}{i+j-1} \frac{(n+i-1)!(n+j-1)!}{[(i-1)!(j-1)!]^2(n-i)!(n-j)!}$$

4. Formulae for the numerical computation of the  $S_n^{ij}$ . It is quite convenient to compute the  $S_n^{ij}$  recursively. We obtain immediately from (3.6)

(4.1) 
$$S_{n+1}^{ij} = \frac{(n+i)(n+j)}{(n+1-i)(n+1-j)} S_n^{ij}$$
 for  $i, j=1,2,...,n$ 

By means of (4.1) it is possible to compute from the elements of  $S_n^{-1}$  the elements of  $S_{n+1}^{-1}$  in the first n rows and columns. We still have to determine the elements of the last column. We can

do this by deriving a relation between the elements of  $S_n^{-1}$  and the coefficients of certain orthogonal polynomials. Consider a polynomial of degree m

(4.2) 
$$P_m(x) = 1 + a_{m,1}x + \dots + a_{m,m}x^m$$
  $(a_{m,m} \neq 0)$ 

such that (4.3

3) 
$$\int_{0}^{1} x^{k} P_{m}(x) dx = 0 \quad \text{for } k = 0, 1, \dots, (m-1)$$

Equation (4.3) determines the polynomials  $P_m(x)$ , they are the Legendre polynomials adapted to the interval  $0 \le x \le 1$ . They are obtained from the Legendre polynomials formed for the interval  $-1 \le x \le +1$  by a linear transformation of the variables. They are discussed and to a certain extent tabulated in [6]. This system of polynomials is orthogonal and we have

(4.4) 
$$\int_{0}^{1} P_{m}(x)P_{n}(x)dx = \begin{cases} 0 \text{ if } m \neq n \\ \frac{1}{2m+1} \text{ if } m = n \end{cases}$$

Moreover,

(4.5) 
$$a_{m,k} = (-1)^k {m \choose k} {m+k}$$

For the proof of (4.4) and (4.5) the reader is referred to W. E. Milne's book [6] chapter IX, Section 69.

We write  $a_{m,0} = 1$  and substitute (4.2) and (4.3) and obtain, considering also (4.4).

$$\sum_{j=0}^{m} \frac{a_{m,j}}{k+j+1} = \frac{1}{(2m+1)a_{m,m}} \delta_{k,m} \quad (k=0,1,\ldots,m),$$

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where  $\delta_{k,m}$  is the Kronecker delta. We rewrite these equations as

(4.6) 
$$\sum_{j=1}^{m+1} \frac{a_{m,j-1}}{k+j-1} = \frac{1}{(2m+1)a_{m,m}} \delta_{k,m+1} \quad (k=1,\ldots,m+1) \quad .$$

Considering the obvious relation

$$\sum_{k=1}^{m+1} \frac{1}{j+k+1} S_{m+1}^{hk} = \delta_{hk}$$

we solve the system (4.6) and obtain

$$a_{m,k-1} = \frac{1}{(2m+1)a_{m,m}} S_{m+1}^{m+1,k}$$

or

(4.7) 
$$S_{m+1}^{m+1,k} = (2m+1) a_{m,k-1} = S_{m+1}^{k,m+1}$$

From (4.7) and (4.5) we see then that for  $k = 1, \dots, (m+1)$ 

(4.8) 
$$S_{m+1}^{m+1,k} = S_{m+1}^{k,m+1} = (-1)^{m+k-1} (2m+1) {\binom{2m}{m}} {\binom{m}{k-1}} {\binom{m+k-1}{k-1}}.$$

Formulae (4.1) and (4.8) can be used to compute systematically tables of the elements of  ${\rm S}_{\rm m}^{-1}$  .

5. An alternate method for computing  $S_m^{-1}$ . We normalize the polynomials (4.2) and obtain

(5.1) 
$$Q_m(x) = \sqrt{2m+1} P_m(x) = \sum_{k=0}^m b_{m,k} x^k$$
 (m=0,1,2,...)

It is seen from (4.5) and (4.4) that

(5.2) 
$$b_{m,k} = (-1)^k \sqrt{2m+1} \binom{m}{k} \binom{m+k}{k}$$

and

(5.3) 
$$\int_0^1 Q_m(x) Q_n(x) dx = \delta_{m,r}$$

We consider the matrix

$$B_{N} = \left\| b_{m,k} \right\|_{m,k=0,\ldots,(N-1)}$$

It follows then from (5.3) that

(5.4)  $B_N S_N B_N' = = I I$ .

Here I denotes the identity matrix and  $B_N^{'}$  the transposed of  $B_{N^{*}}$ . From (5.4) we see that

(5.5) 
$$S_N^{-1} = B_N^{-1}B_N$$
.

If tables of the coefficients of the polynomials  $P_m(x)$  are known it is possible to find  $S_N^{-1}$  from (5.5). On the other hand one could also derive (3.6) from (5.5) and (5.2). 6. Description of Tables and their preparation. Tables are given for the inverse of the matrix  $S_n$  [see equation (3.1)] for n=2,...,10. That is the quantities  $S_n^{ij}$  [see equation (3.2)] are tabulated for  $1 \le i \le j \le n \le 10$ .

Since the matrix  $S_n$  is symmetric its inverse is also, hence in the following tables  $S_n^{ij}$  is given only for  $j \ge i$ .

These tables were obtained by the use of equations (4.1), (4.7) and (4.8).

The tables were checked by using the relationships:

(6.1) 
$$\sum_{i=1}^{n} S_{n}^{ij}/(i+k-1) = \delta_{jk} \quad (j,k = 1,...,n).$$

All of the above relations were checked for j=k, and supplementarily a few others were examined and found correct.

Acknowledgement: The authors wish to thank Mr. Edwin L. Grab who prepared and checked the tables.

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### THE NATIONAL BUREAU OF STANDARDS

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