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NUMERICAL COMPUTATION OF LOW MOMENTS OF ORDER STATISTICS

FROM A NORMAL POPULATION

by

J. Barkley Rosser

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and
University of California at Los Angeles



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U. S. DEPARTMENT OF COMMERCE
NATIONAL BUREAU OF STANDARDS

Numerical Computation of Low Moments of Order Statistics
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1. Summary. The moments considered in this paper are the first eight moments, about the origin, (not including cross moments) of the order statistics for a sample drawn from a normal population (the precise definition is given at the end of Section 3 below). These are explicitly expressed as linear combinations of the moments of the central observations. These, in turn, are tabulated for samples of twenty or fewer, and an asymptotic series is provided which will give their values with great accuracy for samples of more than twenty.

For samples of twenty or fewer, the mean values of the j -th observation (in magnitude) are tabulated. There is also given an asymptotic formula for these means which is more accurate than the classic asymptotic formula. For the means and standard deviations of the extreme observations of a sample, still more accurate asymptotic formulas are furnished.

2. Algebraic preliminaries. Suppose we have quantities $\alpha(m,n)$ satisfying the relations

$$(1) \quad \alpha(m+1,n) + \alpha(m,n+1) = \alpha(m,n)$$

¹The preparation of this paper was sponsored (in part) by the Office of Naval Research.

where m and n are any non-negative integers. Clearly a given linear relationship among the α 's can be derived from (1) if and only if it would become a polynomial identity upon replacing $\alpha(m,n)$ by $x^m(1-x)^n$.

We define

$$(2) \quad P_N(w) = 1 - (N-2)w + \frac{(N-3)(N-4)}{2!} w^2 + \dots + (-1)^k \binom{N-k-1}{k} w^k$$

with

$$k = \left[\frac{N-1}{2} \right] ,$$

Then by induction on N , one can prove for $N \geq 1$,

$$(3) \quad x^N = x P_N(x(1-x)) - x(1-x) P_{N-1}(x(1-x)) .$$

Multiplying by $x^n(1-x)^n$, we see that from (1) can be derived for $N \geq 1$

$$\alpha(n+N, n) = \left\{ \alpha(n+1, n) - (N-2) \alpha(n+2, n+1) + \dots + (-1)^k \binom{N-k-1}{k} \alpha(n+k+1, n+k) \right\} \\ - \left\{ \alpha(n+1, n+1) - (N-3) \alpha(n+2, n+2) + \dots + (-1)^{\gamma} \binom{N-\gamma-2}{\gamma} \alpha(n+\gamma+1, n+\gamma+1) \right\} ,$$

with k as before and

$$\gamma = \left[\frac{N-2}{2} \right] .$$

Thus, by use of (1) each $\alpha(m, n)$ with $m \geq n$ can be expressed as a sum of terms of the form $\alpha(r+1, r)$ and $\alpha(r, r)$.

When there is a symmetry relation

$$\alpha(m, n) = \alpha(n, m) ,$$

then one has by (1) that

$$\alpha(r+1, r) = \frac{1}{2} \alpha(r, r) .$$

In such case, every $\alpha(m, n)$ can be expressed as a sum of terms of the form $\alpha(r, r)$.

When there is an antisymmetry relation

$$\alpha(m, n) = - \alpha(n, m) ,$$

then

$$\alpha(r, r) = 0 ,$$

so that every $\alpha(m, n)$ can be expressed as a sum of terms of the form $\alpha(r+1, r)$.

3. Choice of basic integrals. Let

$$(4) \quad f(x) = (2\pi)^{-1/2} e^{-x^2/2} ,$$

$$(5) \quad F(x) = \int_{-\infty}^x f(y) dy .$$

Then

$$F(-x) = 1 - F(x) ,$$

so that

$$F^m(x) F^n(-x)$$

satisfies (1). So if we multiply by $x^s f(x)$ and integrate, we see that

$$(6) \quad \alpha_s(m,n) = \int_{-\infty}^{\infty} x^s f(x) F^m(x) F^n(-x) dx$$

satisfies (1). Replacing x by $-x$ in the integrand gives further

$$(7) \quad \alpha_{2t}(m,n) = \alpha_{2t}(n,m) ,$$

$$(8) \quad \alpha_{2t+1}(m,n) = - \alpha_{2t+1}(n,m) ,$$

Hence every $\alpha_{2t}(m,n)$ can be expressed as a sum of terms of the form

$\alpha_{2t}(r,r)$ and every $\alpha_{2t+1}(m,n)$ can be expressed as a sum of terms of the form $\alpha_{2t+1}(r+1,r)$. Further, we have

$$\begin{aligned} \frac{d}{dx} \left\{ x^{2t+1} F^{r+1}(x) F^{r+1}(-x) \right\} &= (2t+1) x^{2t} F^{r+1}(x) F^{r+1}(-x) \\ &\quad + (r+1) x^{2t+1} f(x) F^r(x) F^{r+1}(-x) \\ &\quad - (r+1) x^{2t+1} f(x) F^{r+1}(x) F^r(-x) . \end{aligned}$$

Integrating from $-\infty$ to ∞ and using (8) gives

$$(9) \quad \alpha_{2t+1}(r+1,r) = \frac{2t+1}{2(r+1)} \int_{-\infty}^{\infty} x^{2t} F^{r+1}(x) F^{r+1}(-x) dx ,$$

By defining

$$(10) \quad \beta_t(r) = \int_{-\infty}^{\infty} x^{2t} F^r(x) F^r(-x) dx ,$$

we can rewrite (9) as

$$\alpha_{2t+1}(r+1,r) = \frac{2t+1}{2(r+1)} \beta_t(r+1) .$$

Then we see that every $\alpha_s(m,n)$ can be written as a sum of terms of the form $\alpha_{2t}(r,r)$ or $\beta_t(r)$.

Since the s -th moment, about the origin, of the j -th observation (in magnitude) from a sample of n is given by

$$\frac{n!}{(j-1)!(n-j)!} \alpha_s(j-1, n-j)$$

we see that evaluation of these moments is reduced to evaluation of the $\alpha_{2t}(r,r)$ and the $\beta_t(r)$.

4. Numerical determination of the basic integrals. In Table I numerical values of $4^r \alpha_{2t}(r,r)$ are given for $1 \leq t \leq 4$ and $1 \leq r \leq 10$. In Table II numerical values of $4^r \beta_t(r)$ are given for $0 \leq t \leq 3$ and $1 \leq r \leq 10$. These were mainly computed by numerical quadrature, and it is believed that the last digit shown is correct to within one unit.

For small values of r , one can get exact values (as in [3]):

$$\alpha_{2t}(0,0) = 1 \cdot 3 \cdots (2t-1)$$

$$\alpha_0(r,r) = \frac{(r!)^2}{(2r+1)!}$$

$$\alpha_2(1,1) = \frac{1}{6} - \frac{1}{2\pi\sqrt{3}}$$

$$\alpha_4(1,1) = \frac{1}{2} - \frac{13}{6\pi\sqrt{3}}$$

$$\alpha_6(1,1) = \frac{5}{2} - \frac{37}{3\pi\sqrt{3}}$$

$$\alpha_8(1,1) = \frac{35}{2} - \frac{820}{9\pi\sqrt{3}}$$

$$\alpha_2(2,2) = \frac{1}{30} - \frac{1}{4\pi\sqrt{3}} + \frac{3}{2\pi^2\sqrt{3}} \operatorname{arccot} \sqrt{15}$$

$$\alpha_4(2,2) = \frac{13}{3} \alpha_2(2,2) - \frac{2}{45} + \frac{1}{4\pi^2\sqrt{5}}$$

$$\alpha_6(2,2) = \frac{74}{3} \alpha_2(2,2) - \frac{29}{90} + \frac{87}{40 \pi^2 \sqrt{5}}$$

$$\alpha_8(2,2) = \frac{1640}{9} \alpha_2(2,2) - \frac{139}{54} + \frac{2837}{150 \pi^2 \sqrt{5}}$$

$$\beta_0(1) = \frac{1}{\sqrt{\pi}}$$

$$\beta_1(1) = \frac{5}{6 \sqrt{\pi}}$$

$$\beta_2(1) = \frac{43}{20 \sqrt{\pi}}$$

$$\beta_3(1) = \frac{531}{56 \sqrt{\pi}}$$

$$\beta_0(2) = \frac{1}{2 \sqrt{\pi}} - \frac{3}{\pi \sqrt{\pi}} \operatorname{arccot} \sqrt{8}$$

$$\beta_1(2) = \frac{5}{6} \beta_0(2) - \frac{1}{3 \pi \sqrt{2 \pi}}$$

$$\beta_2(2) = \frac{43}{20} \beta_0(2) - \frac{77}{60 \pi \sqrt{2 \pi}}$$

$$\beta_3(2) = \frac{531}{56} \beta_0(2) - \frac{6617}{1008 \pi \sqrt{2 \pi}},$$

We record the values

$$\operatorname{arccot} \sqrt{15} = .25268 \ 02551 \ 42078 \ 65348$$

$$\operatorname{arccot} \sqrt{8} = .33983 \ 69094 \ 54121 \ 93710$$

To compute $\alpha_{2t}(r,r)$ and $\beta_t(r)$ for $r > 10$, we proceed as follows. Clearly for $x \geq 0$,

$$\begin{aligned} f(x)(1 - 2F(x)) &= -\frac{1}{\pi} e^{-x^2/2} \int_0^x e^{-y^2/2} dy \\ &= -\frac{x}{\pi} e^{-x^2/2} \int_0^1 e^{-x^2 w^2/2} dw. \end{aligned}$$

Integrating from 0 to x gives

$$F(x) F(-x) = \frac{1}{\pi} \int_0^1 \frac{e^{-x^2(1+w^2)/2}}{1+w^2} dw .$$

Then for $0 \leq y$,

$$e^{y/2} F(\sqrt{y}) F(-\sqrt{y})$$

is completely monotonic (as defined in [1]). Hence

$$4^m e^{my/2} \left\{ F(\sqrt{y}) F(-\sqrt{y}) \right\}^m$$

is completely monotonic. So by Theorem B of [1], $G_1(y)$ is completely monotonic, where

$$G_1(y) = \frac{1}{y^2} \left\{ 4^m e^{my/2} \left\{ F(\sqrt{y}) F(-\sqrt{y}) \right\}^m - e^{-(4-\pi)my/2\pi} \right\} .$$

We infer

$$0 \leq G_1(y) \leq G_1(0) = \frac{2(\pi-3)}{3\pi^2} m .$$

Multiplying by $y^2 \exp(-my/2)$ and putting $y = x^2$ gives

$$\begin{aligned} 0 &\leq 4^m \left\{ F(x) F(-x) \right\}^m - e^{-2mx^2/\pi} \\ &\leq \frac{2(\pi-3)}{3\pi^2} mx^4 e^{-mx^2/2} . \end{aligned}$$

Putting $m = 1$ gives for $x \geq 0$

$$\left\{ 1 - e^{-2x^2/\pi} - \frac{2(\pi-3)}{3\pi^2} x^4 e^{-x^2/2} \right\}^{1/2}$$

$$\leq \frac{1}{\sqrt{2\pi}} \int_{-x}^x e^{-t^2/2} dt \leq \left\{ 1 - e^{-2x^2/\pi} \right\}^{1/2} .$$

From this, one quickly derives many of the results of [2].

Since $G_1(y)$ is completely monotonic, we can apply Theorem B of [1] again to infer that $G_2(y)$ is completely monotonic, where

$$G_2(y) = \frac{1}{y^2} \left\{ G_1(y) - m \frac{2(\pi-3)}{3\pi^2} \exp \left(- \frac{(4-\pi)m}{2\pi} y - \frac{7\pi^2 - 60\pi + 120}{30\pi(\pi-3)} y \right) \right\} .$$

So

$$0 \leq G_2(y) \leq G_2(0) .$$

But

$$G_2(0) = \frac{2(\pi-3)^2}{9\pi^4} m^2 + \left\{ \frac{3\pi^3 - 56\pi^2 + 280\pi - 420}{105\pi^4} - \frac{(7\pi^2 - 60\pi + 120)^2}{2700\pi^4(\pi-3)} \right\} m$$

$$= 0.000045737m^2 - 0.000012634m .$$

Multiplying by $y^4 \exp(-my/2)$ and putting $y = x^2$ gives

$$0 \leq 4^m \left\{ F(x) F(-x) \right\}^m - e^{-2mx^2/\pi}$$

$$- \frac{2(\pi-3)}{3\pi^2} mx^4 \exp \left(- \frac{2mx^2}{\pi} - \frac{7\pi^2 - 60\pi + 120}{30\pi(\pi-3)} x^2 \right)$$

$$\leq \left\{ 0.00005m^2 - 0.00001m \right\} x^8 e^{-mx^2/2} .$$

Proceeding to make two more uses of Theorem B of [1] gives

$$\begin{aligned} 0 &\leq 4^m \left\{ F(x) F(-x) \right\}^m - \theta(m, x) \\ &\leq \phi(m) x^{16} e^{-mx^2/2}, \end{aligned}$$

where

$$\begin{aligned} \theta(m, x) &= e^{-2mx^2/\pi} \\ &+ K_1 m x^4 \exp \left(-\frac{2mx^2}{\pi} - K_2 x^2 \right) \\ &+ (K_3 m^2 - K_4 m) x^8 \exp \left(-\frac{2mx^2}{\pi} - K_2 x^2 - K_2 \frac{m - K_5}{m - K_6} x^2 \right) \\ &+ \frac{K_7 m^4 - K_8 m^3 + K_9 m^2 - K_{10} m}{m - K_6} x^{12} \exp \left(-\frac{2mx^2}{\pi} - K_2 x^2 \right. \\ &\quad \left. - K_2 \frac{m - K_5}{m - K_6} x^2 - \frac{K_{11} m^4 - K_{12} m^3 + K_{13} m^2 - K_{14} m + K_{15}}{(m - K_6)(K_7 m^3 - K_8 m^2 + K_9 m - K_{10})} x^2 \right), \\ \phi(m) &= \left\{ K_{16} m^{10} - K_{17} m^9 + K_{18} m^8 - K_{19} m^7 + K_{20} m^6 - K_{21} m^5 + K_{22} m^4 - K_{23} m^3 \right. \\ &\quad \left. + K_{24} m^2 - K_{25} m \right\} \left\{ (m - K_6)^3 (K_7 m^3 - K_8 m^2 + K_9 m - K_{10}) \right\}^{-1}, \end{aligned}$$

and

$$\frac{2}{\pi} = 0.63661\ 97723\ 67581\ 34308$$

$$K_1 = \frac{2(\pi - 3)}{3\pi^2} = 9.5642\ 23504\ 51823\ 814 \times 10^{-3}$$

$$K_2 = \frac{7\pi^2 - 60\pi + 120}{30\pi(\pi - 3)} = 4.4337\ 26195\ 26472\ 9694 \times 10^{-2}$$

$$K_3 = 4.5737\ 18562\ 21895\ 644 \times 10^{-5}$$

$$K_4 = 1.2634\ 27027\ 83428\ 618 \times 10^{-5}$$

$$K_5 = 0.71005\ 42616\ 12505\ 98$$

$$K_6 = K_4/K_3 = 0.27623\ 62857\ 80848\ 29$$

$$K_7 = 1.4581\ 35552\ 52753\ 02 \times 10^{-7}$$

$$K_8 = 1.6111\ 59796\ 78083\ 98 \times 10^{-7}$$

$$K_9 = 7.3706\ 39743\ 26849\ 1 \times 10^{-8}$$

$$K_{10} = 1.9600\ 16255\ 51239\ 8 \times 10^{-8}$$

$$K_{11} = 6.4649\ 73795\ 48812\ 1 \times 10^{-9}$$

$$K_{12} = 1.4538\ 54543\ 52510\ 69 \times 10^{-8}$$

$$K_{13} = 1.1327\ 56568\ 89158\ 99 \times 10^{-8}$$

$$K_{14} = 3.1274\ 67461\ 71943\ 6 \times 10^{-9}$$

$$K_{15} = 1.5987\ 53417\ 33771 \times 10^{-10}$$

$$K_{16} = 5.0837\ 65663 \times 10^{-17}$$

$$K_{17} = 1.8256\ 16708 \times 10^{-16}$$

$$K_{18} = 3.1469\ 17320 \times 10^{-16}$$

$$K_{19} = 3.5809\ 66407 \times 10^{-16}$$

$$K_{20} = 2.9053\ 68156 \times 10^{-16}$$

$$K_{21} = 2.6263\ 40177 \times 10^{-16}$$

$$K_{22} = 5.6170\ 32050 \times 10^{-17}$$

$$K_{23} = 1.1854\ 05126 \times 10^{-17}$$

$$K_{24} = 1.3845\ 19506 \times 10^{-18}$$

$$K_{25} = 6.8859\ 48529 \times 10^{-20}$$

We have then

$$0 \leq 4^m \beta_t(m) - \int_{-\infty}^{\infty} x^{2t} \theta(m, x) dx$$

$$\leq \phi(m) \frac{1 \cdot 3 \cdot 5 \cdots (2t+15)}{m^{8+t}} \sqrt{\frac{2\pi}{m}} .$$

However, we can write

$$\int_{-\infty}^{\infty} x^{2t} \theta(m, x) dx$$

in closed form; indeed the leading (and dominant) term is

$$\frac{1 \cdot 3 \cdot 5 \cdots (2t-1) \pi^{t+1}}{(4m)^t \sqrt{2m}} .$$

Also for $m \geq 11$,

$$\phi(m) \leq (4.2) m^4 10^{-10} .$$

Thus for $m \geq 11$, we have quite an accurate approximation to $4^m \beta_t(m)$.

If a more accurate approximation is desired, one can carry out a numerical quadrature of the function

$$x^{2t} \left(4^m \{F(x) F(-x)\}^m - \theta(m,x) \right) .$$

For large m , this function is both small and very smooth, and one can readily get a high accuracy by a simple numerical quadrature.

In a similar fashion, we can get accurate numerical approximations for the $\alpha_{2t}(r,r)$.

5. Estimates for means. In computing moments of arbitrary observations from the moments of central observations by the formulas given in Section 2, it is clear that there will be considerable loss of significant figures for the outlying observations when the sample size is large. We now have sufficiently accurate values for the central observations that this effect is not damaging for samples of reasonable size. Thus, we used this method to compute the means of observations for samples of size twenty or fewer. The results are tabulated in Table III, where $m(j|n)$ denotes the mean of the j -th observation from a sample of size n . We have retained as many decimals as we think are accurate, but do not guarantee the final digit shown. Our table extends the tables of means given in [3] and [4]. Clearly, one could use our estimates of $\beta_0(r)$ to extend the accuracy of Table XX of [5].

Incidentally, in making up such a table, one would start with the $\alpha(r,r)$ and $\alpha(r+1,r)$ and work outward by means of the relation (1). Our explicit formula for $\alpha(n+N,n)$ would be used only to estimate the accumulation of errors, or to compute an isolated value.

Quite clearly, the method outlined above for computing moments is essentially a recursive procedure, in that the computation of means for a sample of a hundred would require the computation of a great many means for observations from smaller samples. Also for large sample sizes, the loss of significant figures begins to be serious for the outlying observations. For these reasons, it seems worthwhile to provide some way of estimating means of observations in a large sample, particularly the outlying observations.

For $m(j|n)$, there is a classical asymptotic estimate, which we shall denote by $\gamma(j|n)$. It is defined by the equation

$$F(\gamma(j|n)) = \frac{j}{n+1} .$$

Some discussion of this estimate is given in [4] and [6]. Values of $\gamma(j|n)$ are tabulated in Table III, from which it is clear that even for a sample of size twenty, the estimate is rather poor, especially for outlying observations.

To obtain a better estimate, we start by inquiring for the value of δ which minimizes

$$(11) \quad \int_{-\infty}^{\infty} (x - \delta)^2 f(x) F^{j-1}(x) F^{n-j}(-x) dx .$$

Clearly (11) can be expanded to

$$\alpha_2(j-1, n-j) - 2\delta \alpha_1(j-1, n-j) + \delta^2 \frac{(j-1)! (n-j)!}{n!} .$$

So (11) is a minimum when

$$\delta = \frac{n!}{(j-1)! (n-j)!} \alpha_1(j-1, n-j) = m(j|n) .$$

That is, $m(j|n)$ is the value of δ which minimizes (11). So we estimate $m(j|n)$ by seeking a value of δ which makes (11) small. To this end, we first observe that

$$(12) \quad f(x) F^{j-1}(x) F^{n-j}(-x)$$

is non-negative, with a single maximum. If (12) were exactly symmetrical, then the x which maximizes (12) would obviously coincide with the δ which minimizes (11). Actually, (12) is reasonably symmetric about its maximum for many values of n and j . Thus one is tempted to use the x at which (12) takes its maximum as an estimate for the δ at which (11) takes its minimum. Accordingly, we define $\delta(j|n)$ to be the value of x at which (12) takes its maximum, and hope that thereby $\delta(j|n)$ will be a good estimate for $m(j|n)$.

We have listed values of $\delta(j|n)$ in Table III. For all entries in Table III, we have

$$\gamma(j|n) \leq \delta(j|n) \leq m(j|n) .$$

One is led to conjecture that this holds for all j and n , so that $\delta(j|n)$ is a better estimate for $m(j|n)$ than $\gamma(j|n)$.

From Table III, one is tempted to make still another conjecture, namely that except for outlying observations, $\delta(j|n)$ lies about halfway between $\gamma(j|n)$ and $m(j|n)$. Thus we are led to suggest

$$2 \delta(j|m) - \gamma(j|n)$$

as an estimate for $m(j|n)$. This estimate is also tabulated in Table III. The four decimal values listed for $2\delta - \gamma$ may be off by as much as three units in the fourth decimal place, since they were computed directly from the four decimal values of $\gamma(j|n)$ and $\delta(j|n)$.

From Table III, it appears that $2\delta - \gamma$ is a fairly successful estimate for $m(j|n)$ except for the outlying observations, for which it does not furnish a good estimate. The reason for this is fairly clear. As $n \rightarrow \infty$, the quantity $\delta(n|n) - \gamma(n|n)$ approaches zero fairly rapidly. Thus γ , δ , and $2\delta - \gamma$ are all about the same for large n . However (see [7]), the asymptotic formula for $m(n|n)$ is

$$(13) \quad \gamma(n|n) + \frac{C}{n f(\gamma(n|n))} ,$$

where C is the Euler constant. For large n , $\gamma(n|n)$, $\delta(n|n)$, and $n f(\gamma(n|n))$ are all nearly equal, and all go very slowly indeed to ∞ as n goes to ∞ . Thus δ , γ , and $2\delta - \gamma$ all become nearly equal before any of them is a good approximation to $m(n|n)$.

For outlying observations near the extreme, the situation is similar, but becomes less severe as one moves further from the extreme (see equation (6) of [7]).

For the extreme observation, an approximation was proposed by Bortkiewicz (see [8]), namely

$$B(n) = \frac{1}{2} \left\{ \delta(n|n) + f(\gamma((n-1)|(n-1))) \sqrt{n(n-1)} \right\} .$$

Sample values of this are listed in Table IV, where for comparison are also listed values of $m(n|n)$. For $n > 20$, the values of $m(n|n)$ are taken from Tippett's table (see [9]).

We wish to propose an alternative formula for $m(n|n)$. Our formula is better than Bortkiewicz's formula for smallish n , say $n \leq 100$, but Bortkiewicz's formula is better for large n , at least as far as $n = 1000$.

One advantage of our formula is that it is less empirical than Bortkiewicz's. Indeed, one may use a similar derivation to get similar estimates for other outlying observations. Furthermore, although our formula appears eventually to become inferior to Bortkiewicz's, the discrepancies are far more uniform, and can be explained. Then by curve fitting methods, one can get a formula which gives a really close fit for $2 \leq n \leq 1000$, and which can be expected to continue to give a close fit for larger n .

We recall that $m(n|n)$ is the value of δ which maximizes

$$\int_{-\infty}^{\infty} (x - \delta)^2 f(x) F^{n-1}(x) dx.$$

Write temporarily $\phi = \delta(n|n)$, so that

$$(14) \quad \phi F(\phi) = (n-1) f(\phi) .$$

Then $m(n|n) - \phi$ is the value of θ which minimizes

$$\int_{-\infty}^{\infty} ((x - \phi) - \theta)^2 f(x) F^{n-1}(x) dx .$$

But this value of θ is given by

$$\theta = \left\{ \int_{-\infty}^{\infty} (x - \phi) f(x) F^{n-1}(x) dx \right\} \left\{ \int_{-\infty}^{\infty} f(x) F^{n-1}(x) dx \right\}^{-1} .$$

To estimate this, we expand the integrands as power series in $x - \phi$.

Using (14), the first few terms are given by

$$f(x) F^{n-1}(x) = \frac{\phi F^n(\phi)}{n-1} \left\{ 1 - \frac{(x-\phi)^2}{2} \left(1 + \frac{\phi^2}{n-1} \right) + \frac{(x-\phi)^3}{6} \phi \left(\frac{\phi^2 n(n+1)}{(n-1)^2} - 1 \right) + \dots \right\}.$$

Thus we are led to propose the approximation

$$f(x) F^{n-1}(x) = \frac{\phi F^n(\phi)}{n-1} \exp \left\{ -\frac{(x-\phi)^2}{2} \left(1 + \frac{\phi^2}{n-1} \right) \right\} \left\{ 1 + \frac{(x-\phi)^3}{6} \phi \left(\frac{\phi^2 n(n+1)}{(n-1)^2} - 1 \right) \right\}.$$

Using this approximation in both the integrands of our formula for θ , we get the approximation

$$\theta = \frac{\phi(\phi^2 n(n+1) - (n-1)^2)}{2(\phi^2 n + n-1)^2}.$$

Thus we are led to the approximation for $m(n|n)$ given by

$$C(n) = \delta(n|n) \left\{ 1 + \frac{(\delta(n|n))^2 n(n+1) - (n-1)^2}{2((\delta(n|n))^2 n + n-1)^2} \right\}.$$

Values of $C(n)$ are listed in Table IV.

For very large n , we have approximately

$$C(n) = \delta(n|n) + \frac{1}{2\delta(n|n)}.$$

Comparison with (13) shows that for large n , $C(n)$ will be less than $m(n|n)$, since $\gamma(n|n)$, $\delta(n|n)$, and $n f(\gamma(n|n))$ are all nearly equal for large n . Indeed, from Table IV, one would conjecture that $C(n) < m(n|n)$ for $n \geq 7$.

In view of (13), one sees that for very large n a better formula than $C(n)$ would be

$$\delta(n|n) + 2C \left\{ C(n) - \delta(n|n) \right\}.$$

Fitting empirically the difference between this and $m(n,n)$, we are led finally to propose

$$D(n) = \delta(n|n) \left\{ 1 + c \frac{(\delta(n|n))^2 n(n+1) - (n-1)^2}{((\delta(n|n))^2 n + n-1)^2} \right\} \\ - \frac{1}{19.541(\delta(n|n))^2 - 66.366 \delta(n|n) + 114.448} .$$

Values of $D(n)$ are listed in Table IV.

The numerical coefficients in the denominator of the second term of $D(n)$ were chosen to give a fit in the region $2 \leq n \leq 100$. Clearly, they could be chosen to give a fit in any other region. However, in view of (13), one might expect high accuracy for $D(n)$ for all values of n .

If one has a good estimate for the mean of the extreme observation, one can derive a good estimate for its standard deviation by using the empirical formula proposed in [10].

Incidentally, by the time one gets to $n = 1000$, the asymptotic formula (13) agrees with $m(n|n)$ to within less than one per cent. One would suspect a similar accuracy for (6) of [7] for other small values of m .

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TABLE I.

Values of $4^r d_{2t}(r,r)$

r	t = 1				t = 2			
1	0.2991	1406	9718	8053	0.4072	7207	9892	6007
2	0.1529	7795	2856	4674	0.1330	4219	4614	9201
3	0.0962	0427	9558	9977	0.0612	8897	0422	4789
4	0.0674	9512	3853	7019	0.0338	9389	5809	0415
5	0.0506	6895	0920	9147	0.0209	8985	6215	1131
6	0.0398	2756	2731	0725	0.0140	3808	4528	5253
7	0.0323	6529	0981	3039	0.0099	2626	8505	3476
8	0.0269	7240	6234	8521	0.0073	2100	7671	5223
9	0.0229	2631	2941	7515	0.0055	8079	0652	7179
10	0.0197	9919	4831	1596	0.0043	6874	1679	6848
	t = 3				t = 4			
1	0.9337	0260	8619	4196	3.0237	4900	0611	929
2	0.1947	6555	2031	6235	0.4028	7649	7414	444
3	0.0656	4164	9173	2218	0.0992	4109	4940	865
4	0.0285	8077	9942	2641	0.0339	8616	3643	807
5	0.0145	8715	6381	2597	0.0142	8346	2227	132
6	0.0082	9464	7482	9277	0.0069	0054	1881	627
7	0.0051	0028	1317	1660	0.0036	8762	1539	726
8	0.0033	2741	7523	9426	0.0021	2706	2546	631
9	0.0022	7385	5368	6564	0.0013	0254	5317	193
10	0.0016	1294	4832	8192	0.0008	3694	8262	323

TABLE II.

Values of $4^r \beta_t(r)$.

r	t = 0				t = 1			
1	2.2567	5833	4191	0251	1.8806	3194	5159	1876
2	1.5840	6070	5464	7751	0.6427	8513	7922	0755
3	1.2898	9973	6330	9072	0.3453	8405	5478	4261
4	1.1155	3389	3536	3525	0.2228	1607	0880	1860
5	0.9969	1887	5707	6363	0.1587	7144	7923	6560
6	0.9095	3967	2162	7550	0.1204	4214	6659	3750
7	0.8417	2526	9829	5399	0.0953	8452	3113	1459
8	0.7871	1902	6160	2866	0.0779	5172	7041	1192
9	0.7419	2432	8542	4813	0.0652	4962	7985	1744
10	0.7037	1536	6288	9335	0.0556	5770	1650	7985
	t = 2				t = 3			
1	4.8520	3041	8510	7041	21.3989	0491	8847	042
2	0.7982	5853	4383	1039	1.6825	9501	5448	810
3	0.2814	2174	8929	4527	0.3873	8273	5294	494
4	0.1349	9461	2115	8106	0.1377	7182	9511	448
5	0.0765	4181	1168	1755	0.0620	3993	4413	791
6	0.0482	0967	3030	9395	0.0324	0067	7160	751
7	0.0326	3892	7355	2676	0.0187	3390	2424	600
8	0.0232	9272	7472	9138	0.0116	6593	9436	320
9	0.0173	0364	7795	6415	0.0076	8672	8404	898
10	0.0132	6717	8669	9766	0.0052	9499	0746	468

TABLE III.

Table of accurate and approximate means.

n	j	m(j n)				$\chi(j n)$	$\delta(j n)$	$2\delta - \chi$
2	2	0.5641	89583	54775	62869	0.4307	0.5061	0.5815
3	3	0.8462	84375	32163	44304	0.6745	0.7653	0.8561
4	4	1.0293	75373	00396	41321	0.8416	0.9359	1.0302
4	3	0.2970	11382	27464	53255	0.2533	0.2775	0.3017
5	5	1.1629	64473	64051	96138	0.9674	1.0615	1.1556
5	4	0.4950	18970	45774	22092	0.4307	0.4654	0.5001
6	6	1.2672	06360	61147	12976	1.0676	1.1602	1.2528
6	5	0.6417	55038	78576	11884	0.5659	0.6058	0.6457
6	4	0.2015	46833	80170	42508	0.1800	0.1919	0.2038
7	7	1.3521	78375	60690	43992	1.1504	1.2412	1.3320
7	6	0.7573	74270	63887	26882	0.6745	0.7171	0.7597
7	5	0.3527	06959	15298	24389	0.3186	0.3371	0.3556
8	8	1.4236	00306	04527	77531	1.2207	1.3095	1.3983
8	7	0.8522	24862	53829	09223	0.7647	0.8087	0.8527
8	6	0.4728	22494	94061	79858	0.4307	0.4533	0.4759
8	5	0.1525	14399	50692	31941	0.1397	0.1467	0.1537
9	9	1.4850	13162	20923	70063	1.2816	1.3684	1.4552
9	8	0.9322	97456	73360	37275	0.8416	0.8863	0.9310
9	7	0.5719	70782	85469	61041	0.5244	0.5495	0.5746
9	6	0.2745	25919	11246	17493	0.2533	0.2649	0.2765
10	10	1.5387	52730	83517	28560	1.3352	1.4202	1.5052
10	9	1.0013	57044	57581	43585	0.9085	0.9534	0.9983
10	8	0.6560	59105	36476	12033	0.6046	0.6313	0.6580
10	7	0.3757	64696	99787	75394	0.3488	0.3633	0.3778
10	6	0.1226	67752	28433	80642	0.1142	0.1188	0.1234
11	11	1.5864	36351	90800	01689	1.3830	1.4662	1.5494
11	10	1.0619	16520	10689	97268	0.9674	1.0124	1.0574
11	9	0.7288	39404	68593	02011	0.6745	0.7023	0.7301
11	8	0.4619	78307	17497	72089	0.4307	0.4474	0.4641
11	7	0.2248	90879	18795	31177	0.2104	0.2183	0.2262
12	12	1.6292	27639	87191	29903	1.4261	1.5076	1.5891
12	11	1.1157	32184	30495	91343	1.0201	1.0648	1.1095
12	10	0.7928	38199	11660	26895	0.7363	0.7649	0.7935
12	9	0.5368	43021	39391	27361	0.5024	0.5206	0.5388
12	8	0.3122	48878	73710	61544	0.2934	0.3036	0.3138
12	7	0.1025	89679	81913	88664	0.0966	0.0999	0.1032
13	13	1.6679	90177	04912	74980	1.4652	1.5452	1.6252

TABLE III (continued)

Table of accurate and approximate means.

n	j	m(j n)				$\gamma(j n)$	$\delta(j n)$	$2\delta - \gamma$
13	12	1.1640	77193	74533	88972	1.0676	1.1120	1.1564
13	11	0.8498	34632	38287	04386	0.7916	0.8207	0.8498
13	10	0.6028	50088	22904	35257	0.5659	0.5852	0.6045
13	9	0.3883	27121	01486	84596	0.3661	0.3780	0.3899
13	8	0.1905	23691	09268	64661	0.1800	0.1857	0.1914
14	14	1.7033	81554	09997	65215	1.5011	1.5796	1.6581
14	13	1.2079	02275	38809	01931	1.1108	1.1548	1.1988
14	12	0.9011	26703	88883	11217	0.8416	0.8709	0.9002
14	11	0.6617	63703	52768	12673	0.6229	0.6430	0.6631
14	10	0.4555	66049	98244	91716	0.4307	0.4440	0.4573
14	9	0.2672	97048	87322	31781	0.2533	0.2609	0.2685
14	8	0.0881	59214	05197	08501	0.0837	0.0862	0.0887
15	15	1.7359	13444	94103	74337	1.5341	1.6113	1.6885
15	14	1.2479	35082	32512	37511	1.1503	1.1940	1.2377
15	13	0.9476	89030	29737	20662	0.8871	0.9166	0.9461
15	12	0.7148	77398	25466	73439	0.6745	0.6952	0.7159
15	11	0.5157	01043	02846	95569	0.4888	0.5030	0.5172
15	10	0.3352	96063	89040	84009	0.3186	0.3278	0.3370
15	9	0.1652	98526	34744	53440	0.1573	0.1617	0.1661
16	16	1.7659	91393	05478	79673	1.5647	1.6406	1.7165
16	15	1.2847	44223	23477	94295	1.1868	1.2301	1.2734
16	14	0.9902	71095	95753	40023	0.9289	0.9584	0.9879
16	13	0.7631	66745	77000	36761	0.7215	0.7426	0.7637
16	12	0.5700	09355	70865	83470	0.5414	0.5564	0.5714
16	11	0.3962	22755	13205	42186	0.3774	0.3875	0.3976
16	10	0.2337	51578	48766	53715	0.2230	0.2289	0.2348
16	9	0.0772	87459	31001	95943	0.0738	0.0757	0.0776
17	17	1.7939	41980	88269	08735	1.5932	1.6679	1.7426
17	16	1.3187	81987	80834	14671	1.2206	1.2635	1.3064
17	15	1.0294	60988	93306	41474	0.9674	0.9969	1.0264
17	14	0.8073	84928	73839	33255	0.7647	0.7861	0.8075
17	13	0.6194	57651	12273	73158	0.5895	0.6051	0.6207
17	12	0.4513	33446	71486	88219	0.4307	0.4417	0.4527
17	11	0.2951	86487	23022	74458	0.2822	0.2892	0.2962
17	10	0.1459	87423	14114	81225	0.1397	0.1431	0.1465

TABLE III (concluded)

Table of accurate and approximate means.

n	j	m(j n)				$\gamma(j n)$	$\delta(j n)$	$2\delta - \gamma$
18	18	1.8200	31878	96872	21046	1.6199	1.6934	1.7669
18	17	1.3504	13713	42015	99452	1.2521	1.2946	1.3371
18	16	1.0657	28182	91379	36422	1.0021	1.0326	1.0631
18	15	0.8481	25019	02941	66731	0.8046	0.8262	0.8478
18	14	0.6647	94612	71981	16086	0.6336	0.6497	0.6658
18	13	0.5015	81550	97034	41543	0.4795	0.4912	0.5029
18	12	0.3508	37238	20391	81572	0.3360	0.3440	0.3520
18	11	0.2077	35307	12871	34707	0.1992	0.2038	0.2084
18	10	0.0688	02568	15669	14373	0.0660	0.0675	0.0690
19	19	1.8444	81511	60382	46581	1.6449	1.7173	1.7897
19	18	1.3799	38491	53687	61420	1.2816	1.3236	1.3656
19	17	1.0994	53099	42807	22722	1.0364	1.0658	1.0952
19	16	0.8858	61961	50430	76155	0.8416	0.8634	0.8852
19	15	0.7066	11484	74857	56393	0.6745	0.6909	0.7073
19	14	0.5477	07371	03927	23228	0.5244	0.5367	0.5490
19	13	0.4016	42274	15433	31226	0.3853	0.3941	0.4029
19	12	0.2637	42690	86034	96451	0.2533	0.2590	0.2647
19	11	0.1307	24879	49771	37309	0.1257	0.1284	0.1311
20	20	1.8674	75059	79832	04847	1.6684	1.7398	1.8112
20	19	1.4076	04095	90840	39521	1.3092	1.3508	1.3924
20	18	1.1309	48052	19312	58507	1.0676	1.0968	1.1260
20	17	0.9209	81700	42610	19940	0.8761	0.8980	0.9199
20	16	0.7453	83005	81713	01013	0.7124	0.7292	0.7460
20	15	0.5902	96921	54291	22532	0.5659	0.5787	0.5915
20	14	0.4483	31753	19744	58187	0.4307	0.4401	0.4495
20	13	0.3149	33241	64569	52584	0.3030	0.3094	0.3158
20	12	0.1869	57364	68233	12250	0.1800	0.1838	0.1876
20	11	0.0619	96286	49429	23493	0.0597	0.0610	0.0623

TABLE IV.

Table of accurate and approximate extreme means.

n	$m(n n)$	B(n)	C(n)	D(n)
2	0.5642	0.5351	0.5654	0.5629
3	0.8463	0.8280	0.8474	0.8467
4	1.0294	1.0183	1.0301	1.0302
5	1.1630	1.1568	1.1633	1.1639
6	1.2672	1.2644	1.2672	1.2681
7	1.3522	1.3518	1.3519	1.3529
8	1.4236	1.4250	1.4231	1.4242
9	1.4850	1.4878	1.4842	1.4855
10	1.5388	1.5426	1.5377	1.5391
11	1.5864	1.5911	1.5852	1.5867
12	1.6292	1.6345	1.6278	1.6294
13	1.6680	1.6738	1.6664	1.6681
14	1.7034	1.7097	1.7017	1.7034
15	1.7359	1.7426	1.7340	1.7358
16	1.7660	1.7729	1.7640	1.7659
17	1.7939	1.8011	1.7918	1.7937
18	1.8200	1.8274	1.8178	1.8198
19	1.8445	1.8520	1.8421	1.8442
20	1.8675	1.8752	1.8650	1.8671
30	2.0428	2.0509	2.0395	2.0422
40	2.1608	2.1687	2.1570	2.1601
50	2.2491	2.2567	2.2449	2.2484
70	2.3774	2.3842	2.3726	2.3767
100	2.5076	2.5134	2.5023	2.5071
200	2.7460	2.7496	2.7398	2.7458
300	2.8778	2.8800	2.8711	2.8777
500	3.0367	3.0374	3.0296	3.0368
700	3.1376	3.1374	3.1301	3.1378
1000	3.2414	3.2403	3.2338	3.2418

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